

# Package ‘bayesm’

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**Depends** R (>= 2.5.1)

**Description** bayesm covers many important models used in marketing and micro-econometrics applications. The package includes: Bayes Regression (univariate or multivariate dep var), Bayes Seemingly Unrelated Regression (SUR), Binary and Ordinal Probit, Multinomial Logit (MNL) and Multinomial Probit (MNP), Multivariate Probit, Negative Binomial (Poisson) Regression, Multivariate Mixtures of Normals (including clustering), Dirichlet Process Prior Density Estimation with normal base, Hierarchical Linear Models with normal prior and covariates, Hierarchical Linear Models with a mixture of normals prior and covariates, Hierarchical Multinomial Logits with a mixture of normals prior and covariates, Hierarchical Multinomial Logits with a Dirichlet Process prior and covariates, Hierarchical Negative Binomial Regression Models, Bayesian analysis of choice-based conjoint data, Bayesian treatment of linear instrumental variables models, and Analysis of Multivariate Ordinal survey data with scale usage heterogeneity (as in Rossi et al, JASA (01)). For further reference, consult our book, Bayesian Statistics and Marketing by Rossi, Allenby and McCulloch.

**License** GPL (version 2 or later)

**URL** <http://www.perossi.org/home/bsm-1>

**Archs** i386, x86\_64

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bank

*Bank Card Conjoint Data of Allenby and Ginter (1995)*

## Description

Data from a conjoint experiment in which two partial profiles of credit cards were presented to 946 respondents. The variable bank\$choiceAtt\$choice indicates which profile was chosen. The profiles are coded as the difference in attribute levels. Thus, a "-1" means the profile coded as a choice of "0" has the attribute. A value of 0 means that the attribute was not present in the comparison.

data on age, income and gender (female=1) are also recorded in bank\$demo

## Usage

```
data(bank)
```

## Format

This R object is a list of two data frames, list(choiceAtt,demo).

List of 2

```
\$ choiceAtt: 'data.frame': 14799 obs. of 16 variables:
```

```
... \$ id : int [1:14799] 1 1 1 1 1 1 1 1 1 1
... \$ choice : int [1:14799] 1 1 1 1 1 1 1 1 0 1
... \$ Med\_FInt : int [1:14799] 1 1 1 0 0 0 0 0 0 0
... \$ Low\_FInt : int [1:14799] 0 0 0 0 0 0 0 0 0 0
... \$ Med\_VInt : int [1:14799] 0 0 0 0 0 0 0 0 0 0
... \$ Rewrd\_2 : int [1:14799] -1 1 0 0 0 0 0 1 -1 0
... \$ Rewrd\_3 : int [1:14799] 0 -1 1 0 0 0 0 0 1 -1
... \$ Rewrd\_4 : int [1:14799] 0 0 -1 0 0 0 0 0 0 1
... \$ Med\_Fee : int [1:14799] 0 0 0 1 1 -1 -1 0 0 0
... \$ Low\_Fee : int [1:14799] 0 0 0 0 0 1 1 0 0 0
... \$ Bank\_B : int [1:14799] 0 0 0 -1 1 -1 1 0 0 0
... \$ Out\_State : int [1:14799] 0 0 0 0 -1 0 -1 0 0 0
... \$ Med\_Rebate : int [1:14799] 0 0 0 0 0 0 0 0 0 0
... \$ High\_Rebate : int [1:14799] 0 0 0 0 0 0 0 0 0 0
... \$ High\_CredLine: int [1:14799] 0 0 0 0 0 0 0 -1 -1 -1
... \$ Long\_Grace : int [1:14799] 0 0 0 0 0 0 0 0 0 0
```

```
\$ demo : 'data.frame': 946 obs. of 4 variables:
```

```
... \$ id : int [1:946] 1 2 3 4 6 7 8 9 10 11
```

```
...$ age : int [1:946] 60 40 75 40 30 30 50 50 50 40
...$ income: int [1:946] 20 40 30 40 30 60 50 100 50 40
...$ gender: int [1:946] 1 1 0 0 0 0 1 0 0 0
```

## Details

Each respondent was presented with between 13 and 17 paired comparisons. Thus, this dataset has a panel structure.

## Source

Allenby and Ginter (1995), "Using Extremes to Design Products and Segment Markets," *JMR*, 392-403.

## References

Appendix A, *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch.  
<http://www.perossi.org/home/bsm-11>

## Examples

```
data(bank)
cat(" table of Binary Dep Var", fill=TRUE)
print(table(bank$choiceAtt[,2]))
cat(" table of Attribute Variables", fill=TRUE)
mat=apply(as.matrix(bank$choiceAtt[,3:16]),2,table)
print(mat)
cat(" means of Demographic Variables", fill=TRUE)
mat=apply(as.matrix(bank$demo[,2:3]),2,mean)
print(mat)

## example of processing for use with rhierBinLogit
##
if(0)
{
choiceAtt=bank$choiceAtt
Z=bank$demo

## center demo data so that mean of random-effects
## distribution can be interpreted as the average respondent

Z[,1]=rep(1,nrow(Z))
Z[,2]=Z[,2]-mean(Z[,2])
Z[,3]=Z[,3]-mean(Z[,3])
Z[,4]=Z[,4]-mean(Z[,4])
Z=as.matrix(Z)

hh=levels(factor(choiceAtt$id))
nhh=length(hh)
lgtdata=NULL
for (i in 1:nhh) {
y=choiceAtt[choiceAtt[,1]==hh[i],2]
nobs=length(y)
X=as.matrix(choiceAtt[choiceAtt[,1]==hh[i],c(3:16)])
lgtdata[[i]]=list(y=y,X=X)
```

```

}

cat("Finished Reading data",fill=TRUE)
fsh()

Data=list(lgtdata=lgtdata,Z=Z)
Mcmc=list(R=10000,sbeta=0.2,keep=20)
set.seed(66)
out=rhierBinLogit(Data=Data,Mcmc=Mcmc)

begin=5000/20
end=10000/20

summary(out$Deltadraw,burnin=begin)
summary(out$Vbetadraw,burnin=begin)

if(0){
## plotting examples

## plot grand means of random effects distribution (first row of Delta)
index=4*c(0:13)+1
matplot(out$Deltadraw[,index],type="l",xlab="Iterations/20",ylab="",
main="Average Respondent Part-Worths")

## plot hierarchical coefs
plot(out$betadraw)

## plot log-likelihood
plot(out$llike,type="l",xlab="Iterations/20",ylab="",main="Log Likelihood")

}
}

```

breg

---

*Posterior Draws from a Univariate Regression with Unit Error Variance*


---

## Description

breg makes one draw from the posterior of a univariate regression (scalar dependent variable) given the error variance = 1.0. A natural conjugate, normal prior is used.

## Usage

```
breg(y, X, betabar, A)
```

## Arguments

y	vector of values of dep variable.
X	n (length(y)) x k Design matrix.
betabar	k x 1 vector. Prior mean of regression coefficients.
A	Prior precision matrix.

## Details

model:  $y = x'\beta + e$ .  $e \sim N(0, 1)$ .

prior:  $\beta \sim N(\text{betabar}, A^{-1})$ .

## Value

k x 1 vector containing a draw from the posterior distribution.

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

In particular, X must be a matrix. If you have a vector for X, coerce it into a matrix with one column

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch.  
<http://www.perossi.org/home/bsm-1>

## Examples

```
##

if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=1000} else {R=10}

## simulate data
set.seed(66)
n=100
X=cbind(rep(1,n),runif(n)); beta=c(1,2)
y=X%%beta+rnorm(n)
##
## set prior
A=diag(c(.05,.05)); betabar=c(0,0)
##
## make draws from posterior
betadraw=matrix(double(R*2),ncol=2)
for (rep in 1:R) {betadraw[rep,]=breg(y,X,betabar,A)}
##
## summarize draws
mat=apply(betadraw,2,quantile,probs=c(.01,.05,.5,.95,.99))
mat=rbind(beta,mat); rownames(mat)[1]="beta"; print(mat)
```

---

cgetC

---

*Obtain A List of Cut-offs for Scale Usage Problems*

---

**Description**

cgetC obtains a list of censoring points, or cut-offs, used in the ordinal multivariate probit model of Rossi et al (2001). This approach uses a quadratic parameterization of the cut-offs. The model is useful for modeling correlated ordinal data on a scale from 1, ..., k with different scale usage patterns.

**Usage**

```
cgetC(e, k)
```

**Arguments**

e	quadratic parameter (>0 and less than 1)
k	items are on a scale from 1, ..., k

**Value**

A vector of k+1 cut-offs.

**Warning**

This is a utility function which implements **no** error-checking.

**Author(s)**

Rob McCulloch and Peter Rossi, Graduate School of Business, University of Chicago. <perossichi@gmail.com>.

**References**

Rossi et al (2001), "Overcoming Scale Usage Heterogeneity," *JASA*96, 20-31.

**See Also**

[rscaleUsage](#)

**Examples**

```
##  
cgetC(.1,10)
```

cheese

*Sliced Cheese Data***Description**

Panel data with sales volume for a package of Borden Sliced Cheese as well as a measure of display activity and price. Weekly data aggregated to the "key" account or retailer/market level.

**Usage**

```
data(cheese)
```

**Format**

A data frame with 5555 observations on the following 4 variables.

RETAILER a list of 88 retailers

VOLUME unit sales

DISP a measure of display activity – per cent ACV on display

PRICE in \\$

**Source**

Boatwright et al (1999), "Account-Level Modeling for Trade Promotion," *JASA* 94, 1063-1073.

**References**

Chapter 3, *Bayesian Statistics and Marketing* by Rossi et al.

<http://www.perossi.org/home/bsm-11>

**Examples**

```
data(cheese)
cat(" Quantiles of the Variables ",fill=TRUE)
mat=apply(as.matrix(cheese[,2:4]),2,quantile)
print(mat)

##
## example of processing for use with rhierLinearModel
##
if(0)
{

retailer=levels(cheese$RETAILER)
nreg=length(retailer)
nvar=3
regdata=NULL
for (reg in 1:nreg) {
y=log(cheese$VOLUME[cheese$RETAILER==retailer[reg]])
iota=c(rep(1,length(y)))
X=cbind(iota,cheese$DISP[cheese$RETAILER==retailer[reg]],
log(cheese$PRICE[cheese$RETAILER==retailer[reg]]))
regdata[[reg]]=list(y=y,X=X)
```



```

}
Z=matrix(c(rep(1,nreg)),ncol=1)
nz=ncol(Z)
##
## run each individual regression and store results
##
lscoef=matrix(double(nreg*nvar),ncol=nvar)
for (reg in 1:nreg) {
  coef=lsfit(regdata[[reg]]$X,regdata[[reg]]$y,intercept=FALSE)$coef
  if (var(regdata[[reg]]$X[,2])==0) { lscoef[reg,1]=coef[1]; lscoef[reg,3]=coef[2]}
  else {lscoef[reg,]=coef }
}

R=2000
Data=list(regdata=regdata,Z=Z)
Mcmc=list(R=R,keep=1)

set.seed(66)
out=rhierLinearModel(Data=Data,Mcmc=Mcmc)

cat("Summary of Delta Draws",fill=TRUE)
summary(out$Deltadraw)
cat("Summary of Vbeta Draws",fill=TRUE)
summary(out$Vbetadraw)

if(0){
#
# plot hier coefs
plot(out$betadraw)
}

}

```

---

clusterMix

---

Cluster Observations Based on Indicator MCMC Draws

---

## Description

clusterMix uses MCMC draws of indicator variables from a normal component mixture model to cluster observations based on a similarity matrix.

## Usage

```
clusterMix(zdraw, cutoff = 0.9, SILENT = FALSE)
```

## Arguments

zdraw	R x nobs array of draws of indicators
cutoff	cutoff probability for similarity (def=.9)
SILENT	logical flag for silent operation (def= FALSE)

## Details

define a similarity matrix,  $\text{Sim}$ ,  $\text{Sim}[i,j]=1$  if observations  $i$  and  $j$  are in same component. Compute the posterior mean of  $\text{Sim}$  over indicator draws.

clustering is achieved by two means:

Method A: Find the indicator draw whose similarity matrix minimizes,  $\text{loss}(E[\text{Sim}]-\text{Sim}(z))$ , where loss is absolute deviation.

Method B: Define a Similarity matrix by setting any element of  $E[\text{Sim}] = 1$  if  $E[\text{Sim}] > \text{cutoff}$ . Compute the clustering scheme associated with this "windsorized" Similarity matrix.

## Value

clustera	indicator function for clustering based on method A above
clusterb	indicator function for clustering based on method B above

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

## Author(s)

Peter Rossi, Graduate School of Business, University of Chicago <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch Chapter 3.

<http://www.perossi.org/home/bsm-1>

## See Also

[rnmixGibbs](#)

## Examples

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0)
{
  ## simulate data from mixture of normals
  n=500
  pvec=c(.5,.5)
  mu1=c(2,2)
  mu2=c(-2,-2)
  Sigma1=matrix(c(1,.5,.5,1),ncol=2)
  Sigma2=matrix(c(1,.5,.5,1),ncol=2)
  comps=NULL
  comps[[1]]=list(mu1,backsolve(chol(Sigma1),diag(2)))
  comps[[2]]=list(mu2,backsolve(chol(Sigma2),diag(2)))
  dm=rmixture(n,pvec,comps)
  ## run MCMC on normal mixture
  R=2000
  Data=list(y=dm$x)
  ncomp=2
}
```

```

Prior=list(ncomp=ncomp,a=c(rep(100,ncomp)))
Mcmc=list(R=R,keep=1)
out=rnmixGibbs(Data=Data,Prior=Prior,Mcmc=Mcmc)
begin=500
end=R
## find clusters
outclusterMix=clusterMix(out$zdraw[begin:end,])
##
## check on clustering versus "truth"
## note: there could be switched labels
##
table(outclusterMix$clustera,dm$z)
table(outclusterMix$clusterb,dm$z)
}
##

```

---

condMom	<i>Computes Conditional Mean/Var of One Element of MVN given All Others</i>
---------	---

---

## Description

condMom compute moments of conditional distribution of ith element of normal given all others.

## Usage

```
condMom(x, mu, sigi, i)
```

## Arguments

x	vector of values to condition on - ith element not used
mu	length(x) mean vector
sigi	length(x) dim inverse of covariance matrix
i	conditional distribution of ith element

## Details

$x \sim MVN(mu, Sigma)$ .

condMom computes moments of  $x_i$  given  $x_{-i}$ .

## Value

a list containing:

cmean	cond mean
cvar	cond variance

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch.  
<http://www.perossi.org/home/bsm-1>

**Examples**

```
##
sig=matrix(c(1,.5,.5,.5,1,.5,.5,.5,1),ncol=3)
sigi=chol2inv(chol(sig))
mu=c(1,2,3)
x=c(1,1,1)
condMom(x,mu,sigi,2)
```

---

createX

---

*Create X Matrix for Use in Multinomial Logit and Probit Routines*


---

**Description**

createX makes up an X matrix in the form expected by Multinomial Logit ([rmnlIndepMetrop](#) and [rhierMnlRwMixture](#)) and Probit ([rmnpGibbs](#) and [rmvpGibbs](#)) routines. Requires an array of alternative specific variables and/or an array of "demographics" or variables constant across alternatives which may vary across choice occasions.

**Usage**

```
createX(p, na, nd, Xa, Xd, INT = TRUE, DIFF = FALSE, base = p)
```

**Arguments**

p	integer - number of choice alternatives
na	integer - number of alternative-specific vars in Xa
nd	integer - number of non-alternative specific vars
Xa	n x p*na matrix of alternative-specific vars
Xd	n x nd matrix of non-alternative specific vars
INT	logical flag for inclusion of intercepts
DIFF	logical flag for differencing wrt to base alternative
base	integer - index of base choice alternative note: na,nd,Xa,Xd can be NULL to indicate lack of Xa or Xd variables.

**Value**

X matrix –  $n \times (p - \text{DIFF}) \times [(\text{INT} + \text{nd}) \times (p - 1) + \text{na}]$  matrix.

**Note**

`rmnpGibbs` assumes that the base alternative is the default.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch.  
<http://www.perossi.org/home/bsm-1>

**See Also**

`rmnlIndepMetrop`, `rmnpGibbs`

**Examples**

```
na=2; nd=1; p=3
vec=c(1,1.5,.5,2,3,1,3,4.5,1.5)
Xa=matrix(vec,byrow=TRUE,ncol=3)
Xa=cbind(Xa,-Xa)
Xd=matrix(c(-1,-2,-3),ncol=1)
createX(p=p,na=na,nd=nd,Xa=Xa,Xd=Xd)
createX(p=p,na=na,nd=nd,Xa=Xa,Xd=Xd,base=1)
createX(p=p,na=na,nd=nd,Xa=Xa,Xd=Xd,DIFF=TRUE)
createX(p=p,na=na,nd=nd,Xa=Xa,Xd=Xd,DIFF=TRUE,base=2)
createX(p=p,na=na,nd=NULL,Xa=Xa,Xd=NULL)
createX(p=p,na=NULL,nd=nd,Xa=NULL,Xd=Xd)
```

---

customerSat

---

*Customer Satisfaction Data*


---

**Description**

Responses to a satisfaction survey for a Yellow Pages advertising product. All responses are on a 10 point scale from 1 to 10 (10 is "Excellent" and 1 is "Poor")

**Usage**

```
data(customerSat)
```

**Format**

A data frame with 1811 observations on the following 10 variables.

- q1 Overall Satisfaction
- q2 Setting Competitive Prices
- q3 Holding Price Increase to a Minimum
- q4 Appropriate Pricing given Volume
- q5 Demonstrating Effectiveness of Purchase

q6 Reach a Large \# of Customers  
 q7 Reach of Advertising  
 q8 Long-term Exposure  
 q9 Distribution  
 q10 Distribution to Right Geographic Areas

### Source

Rossi et al (2001), "Overcoming Scale Usage Heterogeneity," *JASA* 96, 20-31.

### References

Case Study 3, *Bayesian Statistics and Marketing* by Rossi et al.  
<http://www.perossi.org/home/bsm-1>\*

### Examples

```
data(customerSat)
apply(as.matrix(customerSat),2,table)
```

---

detailing

*Physician Detailing Data from Manchanda et al (2004)*

---

### Description

Monthly data on detailing (sales calls) on 1000 physicians. 23 mos of data for each Physician. Includes physician covariates. Dependent Variable (scripts) is the number of new prescriptions ordered by the physician for the drug detailed.

### Usage

```
data(detailing)
```

### Format

This R object is a list of two data frames, list(counts,demo).

List of 2:

```
\$ counts:'data.frame': 23000 obs. of 4 variables:
... \$ id : int [1:23000] 1 1 1 1 1 1 1 1 1 1
... \$ scripts : int [1:23000] 3 12 3 6 5 2 5 1 5 3
... \$ detailing : int [1:23000] 1 1 1 2 1 0 2 2 1 1
... \$ lagged\_scripts: int [1:23000] 4 3 12 3 6 5 2 5 1 5

\$ demo : 'data.frame': 1000 obs. of 4 variables:
... \$ id : int [1:1000] 1 2 3 4 5 6 7 8 9 10
... \$ generalphys : int [1:1000] 1 0 1 1 0 1 1 1 1 1
... \$ specialist: int [1:1000] 0 1 0 0 1 0 0 0 0 0
... \$ mean\_samples: num [1:1000] 0.722 0.491 0.339 3.196 0.348
```

## Details

generalphys is dummy for if doctor is a "general practitioner," specialist is dummy for if the physician is a specialist in the therapeutic class for which the drug is intended, mean\\_samples is the mean number of free drug samples given the doctor over the sample.

## Source

Manchanda, P., P. K. Chintagunta and P. E. Rossi (2004), "Response Modeling with Non-Random Marketing Mix Variables," *Journal of Marketing Research* 41, 467-478.

## Examples

```
data(detailing)
cat(" table of Counts Dep Var", fill=TRUE)
print(table(detailing$counts[,2]))
cat(" means of Demographic Variables", fill=TRUE)
mat=apply(as.matrix(detailing$demo[,2:4]),2,mean)
print(mat)

##
## example of processing for use with rhierNegbinRw
##
if(0)
{
  data(detailing)
  counts = detailing$counts
  Z = detailing$demo

  # Construct the Z matrix
  Z[,1] = 1
  Z[,2]=Z[,2]-mean(Z[,2])
  Z[,3]=Z[,3]-mean(Z[,3])
  Z[,4]=Z[,4]-mean(Z[,4])
  Z=as.matrix(Z)
  id=levels(factor(counts$id))
  nreg=length(id)
  nobs = nrow(counts$id)

  regdata=NULL
  for (i in 1:nreg) {
    X = counts[counts[,1] == id[i],c(3:4)]
    X = cbind(rep(1,nrow(X)),X)
    y = counts[counts[,1] == id[i],2]
    X = as.matrix(X)
    regdata[[i]]=list(X=X, y=y)
  }
  nvar=ncol(X)          # Number of X variables
  nz=ncol(Z)            # Number of Z variables
  rm(detailing,counts)
  cat("Finished Reading data",fill=TRUE)
  fsh()

  Data = list(regdata=regdata, Z=Z)
  deltabar = matrix(rep(0,nvar*nz),nrow=nz)
  Vdelta = 0.01 * diag(nz)
  nu = nvar+3
```

```

V = 0.01*diag(nvar)
a = 0.5
b = 0.1
Prior = list(deltabar=deltabar, Vdelta=Vdelta, nu=nu, V=V, a=a, b=b)

R = 10000
keep =1
s_beta=2.93/sqrt(nvar)
s_alpha=2.93
c=2
Mcmc = list(R=R, keep = keep, s_beta=s_beta, s_alpha=s_alpha, c=c)
out = rhierNegbinRw(Data, Prior, Mcmc)

# Unit level mean beta parameters
Mbeta = matrix(rep(0,nreg*nvar),nrow=nreg)
ndraws = length(out$alphadraw)
for (i in 1:nreg) { Mbeta[i,] = rowSums(out$Betadraw[i, , ])/ndraws }

cat(" Deltadraws ",fill=TRUE)
summary(out$Deltadraw)
cat(" Vbetadraws ",fill=TRUE)
summary(out$Vbetadraw)
cat(" alphadraws ",fill=TRUE)
summary(out$alphadraw)

if(0){
## plotting examples
plot(out$betadraw)
plot(out$alphadraw)
plot(out$Deltadraw)
}
}

```

---

eMixMargDen

---

*Compute Marginal Densities of A Normal Mixture Averaged over MCMC Draws*


---

## Description

eMixMargDen assumes that a multivariate mixture of normals has been fitted via MCMC (using `rnmixGibbs`). For each MCMC draw, the marginal densities for each component in the multivariate mixture are computed on a user-supplied grid and then averaged over draws.

## Usage

```
eMixMargDen(grid, probdraw, compdraw)
```

## Arguments

grid	array of grid points, <code>grid[,i]</code> are ordinates for <i>i</i> th dimension of the density
probdraw	array - each row of which contains a draw of probabilities of mixture comp
compdraw	list of lists of draws of mixture comp moments



**Details**

length(compdraw) is number of MCMC draws.  
 compdraw[[i]] is a list draws of mu and inv Chol root for each of mixture components.  
 compdraw[[i]][[j]] is jth component. compdraw[[i]][[j]]\$mu is mean vector; compdraw[[i]][[j]]\$rooti  
 is the UL decomp of  $\Sigma^{-1}$ .

**Value**

an array of the same dimension as grid with density values.

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type. To avoid errors, call with output from [rnmixGibbs](#).

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch.  
<http://www.perossi.org/home/bsm-1>

**See Also**

[rnmixGibbs](#)

---

 fsh

---

*Flush Console Buffer*


---

**Description**

Flush contents of console buffer. This function only has an effect on the Windows GUI.

**Usage**

```
fsh()
```

**Value**

No value is returned.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

ghkvec

*Compute GHK approximation to Multivariate Normal Integrals***Description**

ghkvec computes the GHK approximation to the integral of a multivariate normal density over a half plane defined by a set of truncation points.

**Usage**

```
ghkvec(L, trunpt, above, r)
```

**Arguments**

L	lower triangular Cholesky root of Covariance matrix
trunpt	vector of truncation points
above	vector of indicators for truncation above(1) or below(0)
r	number of draws to use in GHK

**Value**

approximation to integral

**Note**

ghkvec can accept a vector of truncations and compute more than one integral. That is, length(trunpt)/length(above) number of different integrals, each with the same Sigma and mean 0 but different truncation points. See example below for an example with two integrals at different truncation points.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 2.

<http://www.perossi.org/home/bsm-1>

**Examples**

```
##

Sigma=matrix(c(1,.5,.5,1),ncol=2)
L=t(chol(Sigma))
trunpt=c(0,0,1,1)
above=c(1,1)
ghkvec(L, trunpt, above, 100)
```

llmnl

*Evaluate Log Likelihood for Multinomial Logit Model***Description**

llmnl evaluates log-likelihood for the multinomial logit model.

**Usage**

```
llmnl(beta,y, X)
```

**Arguments**

beta	k x 1 coefficient vector
y	n x 1 vector of obs on y (1,..., p)
X	n*p x k Design matrix (use createX to make)

**Details**

Let  $\mu_{i,j} = X_i\beta$ , then  $Pr(y_i = j) = \exp(\mu_{i,j}) / \sum_k \exp(\mu_{i,k})$ .  
 $X_i$  is the submatrix of X corresponding to the ith observation. X has n\*p rows.  
 Use [createX](#) to create X.

**Value**

value of log-likelihood (sum of log prob of observed multinomial outcomes).

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch.  
<http://www.perossi.org/home/bsm-1>

**See Also**

[createX](#), [rmnlIndepMetrop](#)

**Examples**

```
##
## Not run: ll=llmnl(beta,y,X)
```

llmnp

*Evaluate Log Likelihood for Multinomial Probit Model***Description**

llmnp evaluates the log-likelihood for the multinomial probit model.

**Usage**

```
llmnp(beta, Sigma, X, y, r)
```

**Arguments**

beta	k x 1 vector of coefficients
Sigma	(p-1) x (p-1) Covariance matrix of errors
X	X is n*(p-1) x k array. X is from differenced system.
y	y is vector of n indicators of multinomial response (1, ..., p).
r	number of draws used in GHK

**Details**

X is (p-1)\*n x k matrix. Use [createX](#) with DIFF=TRUE to create X.

Model for each obs:  $w = X\beta + e$ .  $e \sim N(0, \Sigma)$ .

censoring mechanism:

if  $y = j (j < p)$ ,  $w_j > \max(w_{-j})$  and  $w_j > 0$   
 if  $y = p$ ,  $w < 0$

To use GHK, we must transform so that these are rectangular regions e.g. if  $y = 1$ ,  $w_1 > 0$  and  $w_1 - w_{-1} > 0$ .

Define  $A_j$  such that if  $j=1, \dots, p-1$ ,  $A_j w = A_j \mu + A_j e > 0$  is equivalent to  $y = j$ . Thus, if  $y=j$ , we have  $A_j e > -A_j \mu$ . Lower truncation is  $-A_j \mu$  and  $cov = A_j \Sigma A_j'$ . For  $j = p$ ,  $e < -\mu$ .

**Value**

value of log-likelihood (sum of log prob of observed multinomial outcomes).

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapters 2 and 4.

<http://www.perossi.org/home/bsm-1>

## See Also

[createX](#), [rmnpGibbs](#)

## Examples

```
##
## Not run: ll=llmnp(beta,Sigma,X,y,r)
```

---

llnhlogit

---

*Evaluate Log Likelihood for non-homothetic Logit Model*


---

## Description

llmnp evaluates log-likelihood for the Non-homothetic Logit model.

## Usage

```
llnhlogit(theta, choice, lnprices, Xexpend)
```

## Arguments

theta	parameter vector (see details section)
choice	n x 1 vector of choice (1, ..., p)
lnprices	n x p array of log-prices
Xexpend	n x d array of vars predicting expenditure

## Details

Non-homothetic logit model with:  $\ln(\psi_i(U)) = \alpha_i - e^{k_i}U$

Structure of theta vector

alpha: (p x 1) vector of utility intercepts.

k: (p x 1) vector of utility rotation parms.

gamma: (k x 1) – expenditure variable coefs.

tau: (1 x 1) – logit scale parameter.

## Value

value of log-likelihood (sum of log prob of observed multinomial outcomes).

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 4.

<http://www.perossi.org/home/bsm-1>

**See Also**

[simnhlogit](#)

**Examples**

```
##
## Not run: ll=llnhlogit(theta,choice,lnprices,Xexpend)
```

---

IndIChisq

---

*Compute Log of Inverted Chi-Squared Density*


---

**Description**

IndIChisq computes the log of an Inverted Chi-Squared Density.

**Usage**

```
IndIChisq(nu, ssq, x)
```

**Arguments**

nu	d.f. parameter
ssq	scale parameter
x	ordinate for density evaluation

**Details**

$Z = \nu * ssq / \chi_{\nu}^2$ ,  $Z \sim$  Inverted Chi-Squared.

IndIChisq computes the complete log-density, including normalizing constants.

**Value**

log density value

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 2.

<http://www.perossi.org/home/bsm-1>

## See Also

[dchisq](#)

## Examples

```
##
IndIChisq(3,1,2)
```

---

IndIWishart	<i>Compute Log of Inverted Wishart Density</i>
-------------	--

---

## Description

IndIWishart computes the log of an Inverted Wishart density.

## Usage

```
IndIWishart(nu, V, IW)
```

## Arguments

nu	d.f. parameter
V	"location" parameter
IW	ordinate for density evaluation

## Details

$Z \sim \text{Inverted Wishart}(\text{nu}, V)$ .

in this parameterization,  $E[Z] = 1/(\text{nu} - k - 1)V$ ,  $V$  is a  $k \times k$  matrix IndIWishart computes the complete log-density, including normalizing constants.

## Value

log density value

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 2.

<http://www.perossi.org/home/bsm-1>

## See Also

[rwishart](#)

## Examples

```
##
IndIWishart(5,diag(3),(diag(3)+.5))
```

---

IndMvn	<i>Compute Log of Multivariate Normal Density</i>
--------	---

---

## Description

IndMvn computes the log of a Multivariate Normal Density.

## Usage

```
IndMvn(x, mu, rooti)
```

## Arguments

x	density ordinate
mu	mu vector
rooti	inv of Upper Triangular Cholesky root of Sigma

## Details

$$z \sim N(\mu, \Sigma)$$

## Value

log density value

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 2.

<http://www.perossi.org/home/bsm-1>



**See Also**[IndMvst](#)**Examples**

```
##
Sigma=matrix(c(1,.5,.5,1),ncol=2)
IndMvn(x=c(rep(0,2)),mu=c(rep(0,2)),rooti=backsolve(chol(Sigma),diag(2)))
```

IndMvst

*Compute Log of Multivariate Student-t Density***Description**

IndMvst computes the log of a Multivariate Student-t Density.

**Usage**

```
IndMvst(x, nu, mu, rooti,NORMC)
```

**Arguments**

x	density ordinate
nu	d.f. parameter
mu	mu vector
rooti	inv of Cholesky root of Sigma
NORMC	include normalizing constant, def: FALSE

**Details**

$$z \sim MVst(mu, nu, \Sigma)$$
**Value**

log density value

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 2.

<http://www.perossi.org/home/bsm-1>

**See Also**[lndMvn](#)**Examples**

```
##
Sigma=matrix(c(1,.5,.5,1),ncol=2)
lndMvst(x=c(rep(0,2)),nu=4,mu=c(rep(0,2)),rooti=backsolve(chol(Sigma),diag(2)))
```

logMargDenNR

*Compute Log Marginal Density Using Newton-Raftery Approx***Description**

logMargDenNR computes log marginal density using the Newton-Raftery approximation.  
 Note: this approximation can be influenced by outliers in the vector of log-likelihoods. Use with **care**.

**Usage**

```
logMargDenNR(l1)
```

**Arguments**

l1                      vector of log-likelihoods evaluated at length(l1) MCMC draws

**Value**

approximation to log marginal density value.

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 6.  
<http://www.perossi.org/home/bsm-11>

margarine

*Household Panel Data on Margarine Purchases***Description**

Panel data on purchases of margarine by 516 households. Demographic variables are included.

**Usage**

```
data(margarine)
```

**Format**

This is an R object that is a list of two data frames, `list(choicePrice,demos)`

List of 2

`\$ choicePrice`: 'data.frame': 4470 obs. of 12 variables:

```
... \$ hhid : int [1:4470] 2100016 2100016 2100016 2100016
... \$ choice : num [1:4470] 1 1 1 1 1 4 1 1 4 1
... \$ PPk\_Stk : num [1:4470] 0.66 0.63 0.29 0.62 0.5 0.58 0.29
... \$ PBB\_Stk : num [1:4470] 0.67 0.67 0.5 0.61 0.58 0.45 0.51
... \$ PFl\_Stk : num [1:4470] 1.09 0.99 0.99 0.99 0.99 0.99 0.99
... \$ PHse\_Stk: num [1:4470] 0.57 0.57 0.57 0.57 0.45 0.45 0.29
... \$ PGen\_Stk: num [1:4470] 0.36 0.36 0.36 0.36 0.33 0.33 0.33
... \$ PImp\_Stk: num [1:4470] 0.93 1.03 0.69 0.75 0.72 0.72 0.72
... \$ PSS\_Tub : num [1:4470] 0.85 0.85 0.79 0.85 0.85 0.85 0.85
... \$ PPk\_Tub : num [1:4470] 1.09 1.09 1.09 1.09 1.07 1.07 1.07
... \$ PFl\_Tub : num [1:4470] 1.19 1.19 1.19 1.19 1.19 1.19 1.19
... \$ PHse\_Tub: num [1:4470] 0.33 0.37 0.59 0.59 0.59 0.59 0.59
```

Pk is Parkay; BB is BlueBonnett, Fl is Fleischmanns, Hse is house, Gen is generic, Imp is Imperial, SS is Shed Spread. `\_Stk` indicates stick, `\_Tub` indicates Tub form.

`\$ demos`: 'data.frame': 516 obs. of 8 variables:

```
... \$ hhid : num [1:516] 2100016 2100024 2100495 2100560
... \$ Income : num [1:516] 32.5 17.5 37.5 17.5 87.5 12.5
... \$ Fs3\_4 : int [1:516] 0 1 0 0 0 0 0 0 0
... \$ Fs5 : int [1:516] 0 0 0 0 0 0 0 0 1
... \$ Fam\_Size : int [1:516] 2 3 2 1 1 2 2 5 2
... \$ college : int [1:516] 1 1 0 0 1 0 1 0 1
... \$ whtcollar: int [1:516] 0 1 0 1 1 0 0 0 1
... \$ retired : int [1:516] 1 1 1 0 0 1 0 1 0
```

Fs3\\_4 is dummy (family size 3-4). Fs5 is dummy for family size  $\geq 5$ . college,whtcollar,retired are dummies reflecting these statuses.

**Details**

choice is a multinomial indicator of one of the 10 brands (in order listed under format). All prices are in `\$`.

## Source

Allenby and Rossi (1991), "Quality Perceptions and Asymmetric Switching Between Brands," *Marketing Science* 10, 185-205.

## References

Chapter 5, *Bayesian Statistics and Marketing* by Rossi et al.  
<http://www.perossi.org/home/bsm-1>

## Examples

```
data(margarine)
cat(" Table of Choice Variable ",fill=TRUE)
print(table(margarine$choicePrice[,2]))
cat(" Means of Prices",fill=TRUE)
mat=apply(as.matrix(margarine$choicePrice[,3:12]),2,mean)
print(mat)
cat(" Quantiles of Demographic Variables",fill=TRUE)
mat=apply(as.matrix(margarine$demos[,2:8]),2,quantile)
print(mat)

##
## example of processing for use with rhierMnlRwMixture
##
if(0)
{
  select= c(1:5,7) ## select brands
  chPr=as.matrix(margarine$choicePrice)
  ## make sure to log prices
  chPr=cbind(chPr[,1],chPr[,2],log(chPr[,2+select]))
  demos=as.matrix(margarine$demos[,c(1,2,5)])

  ## remove obs for other alts
  chPr=chPr[chPr[,2] <= 7,]
  chPr=chPr[chPr[,2] != 6,]

  ## recode choice
  chPr[chPr[,2] == 7,2]=6

  hhid1=levels(as.factor(chPr[,1]))
  lgtdata=NULL
  nlgt=length(hhid1)
  p=length(select) ## number of choice alts
  ind=1
  for (i in 1:nlgt) {
    nobs=sum(chPr[,1]==hhid1[i])
    if(nobs >=5) {
      data=chPr[chPr[,1]==hhid1[i],]
      y=data[,2]
      names(y)=NULL
      X=createX(p=p,na=1,Xa=data[,3:8],nd=NULL,Xd=NULL,INT=TRUE,base=1)
      lgtdata[[ind]]=list(y=y,X=X,hhid=hhid1[i]); ind=ind+1
    }
  }
  nlgt=length(lgtdata)
  ##
```

```

## now extract demos corresponding to hhs in lgtdata
##
Z=NULL
nlgt=length(lgtdata)
for(i in 1:nlgt){
  Z=rbind(Z,demos[demos[,1]==lgtdata[[i]]$hhid,2:3])
}
##
## take log of income and family size and demean
##
Z=log(Z)
Z[,1]=Z[,1]-mean(Z[,1])
Z[,2]=Z[,2]-mean(Z[,2])

keep=5
R=20000
mcmc1=list(keep=keep,R=R)
out=rhierMnlRwMixture(Data=list(p=p,lgtdata=lgtdata,Z=Z),Prior=list(ncomp=1),Mcmc=mcmc1)

summary(out$Deltadraw)
summary(out$nmix)

if(0){
  ## plotting examples
  plot(out$nmix)
  plot(out$Deltadraw)}
}

```

---

mixDen

---

*Compute Marginal Density for Multivariate Normal Mixture*


---

## Description

mixDen computes the marginal density for each component of a normal mixture at each of the points on a user-specified grid.

## Usage

```
mixDen(x, pvec, comps)
```

## Arguments

x	array - ith column gives grid points for ith variable
pvec	vector of mixture component probabilities
comps	list of lists of components for normal mixture

## Details

length(comps) is the number of mixture components. comps[[j]] is a list of parameters of the jth component. comps[[j]]\$mu is mean vector; comps[[j]]\$rooti is the UL decomp of  $\Sigma^{-1}$ .

**Value**

an array of the same dimension as grid with density values.

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 3.

<http://www.perossi.org/home/bsm-1>

**See Also**

[rnmixGibbs](#)

**Examples**

```
## Not run:
##
## see examples in rnmixGibbs documentation
##

## End(Not run)
```

---

mixDenBi

---

*Compute Bivariate Marginal Density for a Normal Mixture*


---

**Description**

mixDenBi computes the implied bivariate marginal density from a mixture of normals with specified mixture probabilities and component parameters.

**Usage**

```
mixDenBi(i, j, xi, xj, pvec, comps)
```

**Arguments**

i	index of first variable
j	index of second variable
xi	grid of values of first variable
xj	grid of values of second variable
pvec	normal mixture probabilities
comps	list of lists of components

## Details

length(comps) is the number of mixture components. comps[[j]] is a list of parameters of the jth component. comps[[j]]\$mu is mean vector; comps[[j]]\$rooti is the UL decomp of  $\Sigma^{-1}$ .

## Value

an array (length(xi)=length(xj) x 2) with density value

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 3.

<http://www.perossi.org/home/bsm-1>

## See Also

[rnmixGibbs](#), [mixDen](#)

## Examples

```
## Not run:
##
## see examples in rnmixGibbs documentation
##

## End(Not run)
```

---

mnHess

*Computes -Expected Hessian for Multinomial Logit*

---

## Description

mnHess computes -Expected[Hessian] for Multinomial Logit Model

## Usage

```
mnHess(beta,y, X)
```

## Arguments

beta	k x 1 vector of coefficients
y	n x 1 vector of choices, (1, ...,p)
X	n*p x k Design matrix

**Details**

See [llmnl](#) for information on structure of X array. Use [createX](#) to make X.

**Value**

k x k matrix

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 3.

<http://www.perossi.org/home/bsm-11>

**See Also**

[llmnl](#), [createX](#), [rmnlIndepMetrop](#)

**Examples**

```
##
## Not run: mnlHess(beta,y,X)
```

---

mnpProb

---

*Compute MNP Probabilities*


---

**Description**

mnpProb computes MNP probabilities for a given X matrix corresponding to one observation. This function can be used with output from rmnpGibbs to simulate the posterior distribution of market shares or fitted probabilities.

**Usage**

```
mnpProb(beta, Sigma, X, r)
```

**Arguments**

beta	MNP coefficients
Sigma	Covariance matrix of latents
X	X array for one observation – use createX to make
r	number of draws used in GHK (def: 100)



**Details**

see [rmnpGibbs](#) for definition of the model and the interpretation of the beta, Sigma parameters. Uses the GHK method to compute choice probabilities. To simulate a distribution of probabilities, loop over the beta, Sigma draws from rmnpGibbs output.

**Value**

p x 1 vector of choice probabilities

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapters 2 and 4.  
<http://www.perossi.org/home/bsm-1>

**See Also**

[rmnpGibbs](#), [createX](#)

**Examples**

```
##
## example of computing MNP probabilities
## here I'm thinking of Xa as having the prices of each of the 3 alternatives
Xa=matrix(c(1,.5,1.5),nrow=1)
X=createX(p=3,na=1,nd=NULL,Xa=Xa,Xd=NULL,DIFF=TRUE)
beta=c(1,-1,-2) ## beta contains two intercepts and the price coefficient
Sigma=matrix(c(1,.5,.5,1),ncol=2)
mnpProb(beta,Sigma,X)
```

---

momMix

---

*Compute Posterior Expectation of Normal Mixture Model Moments*


---

**Description**

momMix averages the moments of a normal mixture model over MCMC draws.

**Usage**

```
momMix(probdraw, compdraw)
```

**Arguments**

probdraw	R x ncomp list of draws of mixture probs
compdraw	list of length R of draws of mixture component moments

## Details

R is the number of MCMC draws in argument list above.  
 ncomp is the number of mixture components fitted.  
 compdraw is a list of lists of lists with mixture components.  
 compdraw[[i]] is ith draw.  
 compdraw[[i]][[j]][[1]] is the mean parameter vector for the jth component, ith MCMC draw.  
 compdraw[[i]][[j]][[2]] is the UL decomposition of  $\text{Sigma}^{-1}$  for the jth component, ith MCMC draw.

## Value

a list of the following items ...

mu	Posterior Expectation of Mean
sigma	Posterior Expectation of Covariance Matrix
sd	Posterior Expectation of Vector of Standard Deviations
corr	Posterior Expectation of Correlation Matrix

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 5.  
<http://www.perossi.org/home/bsm-1>

## See Also

[rmixGibbs](#)

---

nmat	<i>Convert Covariance Matrix to a Correlation Matrix</i>
------	--

---

## Description

nmat converts a covariance matrix (stored as a vector, col by col) to a correlation matrix (also stored as a vector).

## Usage

```
nmat(vec)
```

## Arguments

vec	k x k Cov matrix stored as a k*k x 1 vector (col by col)
-----	--

**Details**

This routine is often used with `apply` to convert an  $R \times (k \times k)$  array of covariance MCMC draws to correlations. As in `corrdraws=apply(vardraws,1,nmat)`

**Value**

$k \times k \times 1$  vector with correlation matrix

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**Examples**

```
##
set.seed(66)
X=matrix(rnorm(200,4),ncol=2)
Varmat=var(X)
nmat(as.vector(Varmat))
```

---

numEff	<i>Compute Numerical Standard Error and Relative Numerical Efficiency</i>
--------	---

---

**Description**

numEff computes the numerical standard error for the mean of a vector of draws as well as the relative numerical efficiency (ratio of variance of mean of this time series process relative to iid sequence).

**Usage**

```
numEff(x, m = as.integer(min(length(x), (100/sqrt(5000)) * sqrt(length(x)))))
```

**Arguments**

x	R x 1 vector of draws
m	number of lags for autocorrelations

**Details**

default for number of lags is chosen so that if  $R = 5000$ ,  $m = 100$  and increases as the  $\sqrt{R}$ .

**Value**

stderr	standard error of the mean of x
f	variance ratio (relative numerical efficiency)

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 3.

<http://www.perossi.org/home/bsm-1>

**Examples**

```
numEff(rnorm(1000),m=20)
numEff(rnorm(1000))
```

---

orangeJuice

---

Store-level Panel Data on Orange Juice Sales

---

**Description**

yx, weekly sales of refrigerated orange juice at 83 stores.  
 storedemo, contains demographic information on those stores.

**Usage**

```
data(orangeJuice)
```

**Format**

This R object is a list of two data frames, list(yx,storedemo).

List of 2

\\$ yx : 'data.frame': 106139 obs. of 19 variables:

... \\$ store : int [1:106139] 2 2 2 2 2 2 2 2 2 2

... \\$ brand : int [1:106139] 1 1 1 1 1 1 1 1 1 1

... \\$ week : int [1:106139] 40 46 47 48 50 51 52 53 54 57

... \\$ logmove : num [1:106139] 9.02 8.72 8.25 8.99 9.09

... \\$ constant: int [1:106139] 1 1 1 1 1 1 1 1 1 1

... \\$ price1 : num [1:106139] 0.0605 0.0605 0.0605 0.0605 0.0605

... \\$ price2 : num [1:106139] 0.0605 0.0603 0.0603 0.0603 0.0603

... \\$ price3 : num [1:106139] 0.0420 0.0452 0.0452 0.0498 0.0436

... \\$ price4 : num [1:106139] 0.0295 0.0467 0.0467 0.0373 0.0311

... \\$ price5 : num [1:106139] 0.0495 0.0495 0.0373 0.0495 0.0495

... \\$ price6 : num [1:106139] 0.0530 0.0478 0.0530 0.0530 0.0530

... \\$ price7 : num [1:106139] 0.0389 0.0458 0.0458 0.0458 0.0466

... \\$ price8 : num [1:106139] 0.0414 0.0280 0.0414 0.0414 0.0414

```

...$ price9 : num [1:106139] 0.0289 0.0430 0.0481 0.0423 0.0423
...$ price10 : num [1:106139] 0.0248 0.0420 0.0327 0.0327 0.0327
...$ price11 : num [1:106139] 0.0390 0.0390 0.0390 0.0390 0.0382
...$ deal : int [1:106139] 1 0 0 0 0 0 1 1 1 1
...$ feat : num [1:106139] 0 0 0 0 0 0 0 0 0 0
...$ profit : num [1:106139] 38.0 30.1 30.0 29.9 29.9

```

1 Tropicana Premium 64 oz; 2 Tropicana Premium 96 oz; 3 Florida's Natural 64 oz;  
 4 Tropicana 64 oz; 5 Minute Maid 64 oz; 6 Minute Maid 96 oz;  
 7 Citrus Hill 64 oz; 8 Tree Fresh 64 oz; 9 Florida Gold 64 oz;  
 10 Dominicks 64 oz; 11 Dominicks 128 oz.

```

\$.storedemo:'data.frame': 83 obs. of 12 variables:
...$ STORE : int [1:83] 2 5 8 9 12 14 18 21 28 32
...$ AGE60 : num [1:83] 0.233 0.117 0.252 0.269 0.178
...$ EDUC : num [1:83] 0.2489 0.3212 0.0952 0.2222 0.2534
...$ ETHNIC : num [1:83] 0.1143 0.0539 0.0352 0.0326 0.3807
...$ INCOME : num [1:83] 10.6 10.9 10.6 10.8 10.0
...$ HHLARGE : num [1:83] 0.1040 0.1031 0.1317 0.0968 0.0572
...$ WORKWOM : num [1:83] 0.304 0.411 0.283 0.359 0.391
...$ HVAL150 : num [1:83] 0.4639 0.5359 0.0542 0.5057 0.3866
...$ SSTRDIST: num [1:83] 2.11 3.80 2.64 1.10 9.20
...$ SSTRVOL : num [1:83] 1.143 0.682 1.500 0.667 1.111
...$ CPDIST5 : num [1:83] 1.93 1.60 2.91 1.82 0.84
...$ CPWVOL5 : num [1:83] 0.377 0.736 0.641 0.441 0.106

```

## Details

store store number  
 brand brand indicator  
 week week number  
 logmove log of the number of units sold  
 constant a vector of 1  
 price1 price of brand 1  
 deal in-store coupon activity  
 feature feature advertisement  
 STORE store number  
 AGE60 percentage of the population that is aged 60 or older  
 EDUC percentage of the population that has a college degree  
 ETHNIC percent of the population that is black or Hispanic  
 INCOME median income  
 HHLARGE percentage of households with 5 or more persons  
 WORKWOM percentage of women with full-time jobs  
 HVAL150 percentage of households worth more than \$150,000  
 SSTRDIST distance to the nearest warehouse store  
 SSTRVOL ratio of sales of this store to the nearest warehouse store  
 CPDIST5 average distance in miles to the nearest 5 supermarkets  
 CPWVOL5 ratio of sales of this store to the average of the nearest five stores

## Source

Alan L. Montgomery (1997), "Creating Micro-Marketing Pricing Strategies Using Supermarket Scanner Data," *Marketing Science* 16(4) 315-337.

## References

Chapter 5, *Bayesian Statistics and Marketing* by Rossi et al.  
<http://www.perossi.org/home/bsm-1>

## Examples

```
## Example
## load data
data(orangeJuice)

## print some quantiles of yx data
cat("Quantiles of the Variables in yx data",fill=TRUE)
mat=apply(as.matrix(orangeJuice$yx),2,quantile)
print(mat)

## print some quantiles of storedemo data
cat("Quantiles of the Variables in storedemo data",fill=TRUE)
mat=apply(as.matrix(orangeJuice$storedemo),2,quantile)
print(mat)

## Example 2 processing for use with rhierLinearModel
##
##
if(0)
{

## select brand 1 for analysis
brand1=orangeJuice$yx[(orangeJuice$yx$brand==1),]

store = sort(unique(brand1$store))
nreg = length(store)
nvar=14

regdata=NULL
for (reg in 1:nreg) {
  y=brand1$logmove[brand1$store==store[reg]]
  iota=c(rep(1,length(y)))
  X=cbind(iota,log(brand1$price1[brand1$store==store[reg]]),
          log(brand1$price2[brand1$store==store[reg]]),
          log(brand1$price3[brand1$store==store[reg]]),
          log(brand1$price4[brand1$store==store[reg]]),
          log(brand1$price5[brand1$store==store[reg]]),
          log(brand1$price6[brand1$store==store[reg]]),
          log(brand1$price7[brand1$store==store[reg]]),
          log(brand1$price8[brand1$store==store[reg]]),
          log(brand1$price9[brand1$store==store[reg]]),
          log(brand1$price10[brand1$store==store[reg]]),
          log(brand1$price11[brand1$store==store[reg]]),
          brand1$deal[brand1$store==store[reg]],
```

```

        brand1$feat[brand1$store==store[reg]]
    regdata[[reg]]=list(y=y,X=X)
}

## storedemo is standardized to zero mean.

Z=as.matrix(orangeJuice$storedemo[,2:12])
dmean=apply(Z,2,mean)
for (s in 1:nreg){
    Z[s,]=Z[s,]-dmean
}
iotaz=c(rep(1,nrow(Z)))
Z=cbind(iotaz,Z)
nz=ncol(Z)

Data=list(regdata=regdata,Z=Z)
Mcmc=list(R=R,keep=1)

out=rhierLinearModel(Data=Data,Mcmc=Mcmc)

summary(out$Deltadraw)
summary(out$Vbetadraw)

if(0){
## plotting examples
plot(out$betadraw)
}
}

```

---

plot.bayesm.hcoef

*Plot Method for Hierarchical Model Coefs*


---

## Description

plot.bayesm.hcoef is an S3 method to plot 3 dim arrays of hierarchical coefficients. Arrays are of class bayesm.hcoef with dimensions: cross-sectional unit x coef x MCMC draw.

## Usage

```
## S3 method for class 'bayesm.hcoef'
plot(x,names,burnin,...)
```

## Arguments

x	An object of S3 class, bayesm.hcoef
names	a list of names for the variables in the hierarchical model
burnin	no draws to burnin, def: .1*R
...	standard graphics parameters

## Details

Typically, `plot.bayesm.hcoef` will be invoked by a call to the generic `plot(object)` where `object` is of class `bayesm.hcoef`. All of the `bayesm` hierarchical routines return draws of hierarchical coefficients in this class (see example below). One can also simply invoke `plot.bayesm.hcoef` on any valid 3-dim array as in `plot.bayesm.hcoef(betadraws)`

`plot.bayesm.hcoef` is also exported for use as a standard function, as in `plot.bayesm.hcoef(array)`.

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## See Also

[rhierMnlRwMixture](#), [rhierLinearModel](#), [rhierLinearMixture](#), [rhierNegbinRw](#)

## Examples

```
##
## not run
# out=rhierLinearModel(Data,Prior,Mcmc)
# plot(out$betadraws)
#
```

---

plot.bayesm.mat

*Plot Method for Arrays of MCMC Draws*

---

## Description

`plot.bayesm.mat` is an S3 method to plot arrays of MCMC draws. The columns in the array correspond to parameters and the rows to MCMC draws.

## Usage

```
## S3 method for class 'bayesm.mat'
plot(x,names,burnin,tvalues,TRACEPLOT,DEN,INT,CHECK_NDRAWS, ...)
```

## Arguments

<code>x</code>	An object of either S3 class, <code>bayesm.mat</code> , or S3 class, <code>mcmc</code>
<code>names</code>	optional character vector of names for coefficients
<code>burnin</code>	number of draws to discard for burn-in, def: <code>.1*nrow(X)</code>
<code>tvalues</code>	vector of true values
<code>TRACEPLOT</code>	logical, TRUE provide sequence plots of draws and acfs, def: TRUE
<code>DEN</code>	logical, TRUE use density scale on histograms, def: TRUE
<code>INT</code>	logical, TRUE put various intervals and points on graph, def: TRUE
<code>CHECK_NDRAWS</code>	logical, TRUE check that there are at least 100 draws, def: TRUE
<code>...</code>	standard graphics parameters



## Details

Typically, `plot.bayesm.mat` will be invoked by a call to the generic plot function as in `plot(object)` where `object` is of class `bayesm.mat`. All of the `bayesm` MCMC routines return draws in this class (see example below). One can also simply invoke `plot.bayesm.mat` on any valid 2-dim array as in `plot.bayesm.mat(betadraws)`.

`plot.bayesm.mat` paints (by default) on the histogram:

green "[" delimiting 95% Bayesian Credibility Interval  
 yellow "(" showing +/- 2 numerical standard errors  
 red "|" showing posterior mean

`plot.bayesm.mat` is also exported for use as a standard function, as in `plot.bayesm.mat(matrix)`

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## Examples

```
##
## not run
# out=runiregGibbs(Data,Prior,Mcmc)
# plot(out$betadraw)
#
```

---

plot.bayesm.nmix

*Plot Method for MCMC Draws of Normal Mixtures*

---

## Description

`plot.bayesm.nmix` is an S3 method to plot aspects of the fitted density from a list of MCMC draws of normal mixture components. Plots of marginal univariate and bivariate densities are produced.

## Usage

```
## S3 method for class 'bayesm.nmix'
plot(x, names, burnin, Grid, bi.sel, nstd, marg, Data, ngrid, ndraw, ...)
```

## Arguments

<code>x</code>	An object of S3 class <code>bayesm.nmix</code>
<code>names</code>	optional character vector of names for each of the dimensions
<code>burnin</code>	number of draws to discard for burn-in, def: <code>.1*nrow(X)</code>
<code>Grid</code>	matrix of grid points for densities, def: mean +/- nstd std deviations (if <code>Data</code> no supplied), range of <code>Data</code> if supplied)
<code>bi.sel</code>	list of vectors, each giving pairs for bivariate distributions, def: <code>list(c(1,2))</code>
<code>nstd</code>	number of standard deviations for default <code>Grid</code> , def: 2
<code>marg</code>	logical, if TRUE display marginals, def: TRUE

Data	matrix of data points, used to paint histograms on marginals and for grid
ngrid	number of grid points for density estimates, def:50
ndraw	number of draws to average Mcmc estimates over, def:200
...	standard graphics parameters

### Details

Typically, `plot.bayesm.nmix` will be invoked by a call to the generic plot function as in `plot(object)` where `object` is of class `bayesm.nmix`. These objects are lists of three components. The first component is an array of draws of mixture component probabilities. The second component is not used. The third is a lists of lists of lists with draws of each of the normal components.

`plot.bayesm.nmix` can also be used as a standard function, as in `plot.bayesm.nmix(list)`.

### Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

### See Also

[rnmixGibbs](#), [rhierMnlRwMixture](#), [rhierLinearMixture](#), [rDPGibbs](#)

### Examples

```
##
## not run
# out=rnmixGibbs(Data,Prior,Mcmc)
# plot(out,bi.sel=list(c(1,2),c(3,4),c(1,3)))
#       # plot bivariate distributions for dimension 1,2; 3,4; and 1,3
#
```

---

rbiNormGibbs

---

*Illustrate Bivariate Normal Gibbs Sampler*


---

### Description

`rbiNormGibbs` implements a Gibbs Sampler for the bivariate normal distribution. Intermediate moves are shown and the output is contrasted with the iid sampler. <sup>i</sup> This function is designed for illustrative/teaching purposes.

### Usage

```
rbiNormGibbs(initx = 2, inity = -2, rho, burnin = 100, R = 500)
```

### Arguments

initx	initial value of parameter on x axis (def: 2)
inity	initial value of parameter on y axis (def: -2)
rho	correlation for bivariate normals
burnin	burn-in number of draws (def:100)
R	number of MCMC draws (def:500)

**Details**

$(\theta_1, \theta_2) \sim N((0,0), \Sigma = \text{matrix}(c(1, \rho, \rho, 1), \text{ncol}=2))$

**Value**

R x 2 array of draws

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapters 2 and 3.

<http://www.perossi.org/home/bsm-1>

**Examples**

```
##
## Not run: out=rbiNormGibbs(rho=.95)
```

---

rbprobitGibbs

*Gibbs Sampler (Albert and Chib) for Binary Probit*


---

**Description**

rbprobitGibbs implements the Albert and Chib Gibbs Sampler for the binary probit model.

**Usage**

```
rbprobitGibbs(Data, Prior, Mcmc)
```

**Arguments**

Data	list(X,y)
Prior	list(betabar,A)
Mcmc	list(R,keep)

**Details**

Model:  $z = X\beta + e$ .  $e \sim N(0, I)$ .  $y=1$ , if  $z > 0$ .

Prior:  $\beta \sim N(\text{betabar}, A^{-1})$ .

List arguments contain

X Design Matrix

y n x 1 vector of observations, (0 or 1)

betabar k x 1 prior mean (def: 0)

A k x k prior precision matrix (def: .01I)

R number of MCMC draws

keep thinning parameter - keep every keepth draw (def: 1)

**Value**

betadraw            R/keep x k array of betadraws

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 3.

<http://www.perossi.org/home/bsm-1>

**See Also**

[rmnpGibbs](#)

**Examples**

```
##
## rbprobitGibbs example
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

set.seed(66)
simbprobit=
function(X,beta) {
  ## function to simulate from binary probit including x variable
  y=ifelse((X*beta+rnorm(nrow(X)))<0,0,1)
  list(X=X,y=y,beta=beta)
}

nobs=200
X=cbind(rep(1,nobs),runif(nobs),runif(nobs))
beta=c(0,1,-1)
nvar=ncol(X)
simout=simbprobit(X,beta)

Data1=list(X=simout$X,y=simout$y)
Mcmc1=list(R=R,keep=1)

out=rbprobitGibbs(Data=Data1,Mcmc=Mcmc1)

summary(out$betadraw,tvalues=beta)

if(0){
  ## plotting example
  plot(out$betadraw,tvalues=beta)
}
```

---

rdirichlet	<i>Draw From Dirichlet Distribution</i>
------------	---

---

**Description**

rdirichlet draws from Dirichlet

**Usage**

```
rdirichlet(alpha)
```

**Arguments**

alpha                vector of Dirichlet parms (must be > 0)

**Value**

Vector of draws from Dirichlet

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 2.

<http://www.perossi.org/home/bsm-1>

**Examples**

```
##  
set.seed(66)  
rdirichlet(c(rep(3,5)))
```

## Description

rDPGibbs implements a Gibbs Sampler to draw from the posterior for a normal mixture problem with a Dirichlet Process prior. A natural conjugate base prior is used along with priors on the hyper parameters of this distribution. One interpretation of this model is as a normal mixture with a random number of components that can grow with the sample size.

## Usage

```
rDPGibbs(Prior, Data, Mcmc)
```

## Arguments

Prior	list(Prioralpha,lambda\_hyper)
Data	list(y)
Mcmc	list(R,keep,maxuniq,SCALE,gridsize)

## Details

Model:

$$y_i \sim N(\mu_i, \Sigma_i).$$

Priors:

$$\theta_i = (\mu_i, \Sigma_i) \sim DP(G_0(\lambda), \alpha)$$

$G_0(\lambda)$  :

$$\mu_i | \Sigma_i \sim N(0, \Sigma_i(x)a^{-1})$$

$$\Sigma_i \sim IW(nu, nu * v * I)$$

$\lambda(a, nu, v)$  :

$$a \sim \text{uniform on grid}[alim[1],alimb[2]]$$

$$nu \sim \text{uniform on grid}[\dim(data)-1 + \exp(nulim[1]),\dim(data)-1 + \exp(nulim[2])]$$

$$v \sim \text{uniform on grid}[vlim[1],vlim[2]]$$

$$\alpha \sim (1 - (\alpha - \alpha_{min})/(\alpha_{max} - \alpha_{min}))^{power}$$

alpha= alphamin then expected number of components = Istarmin

alpha= alphamax then expected number of components = Istarmax

list arguments

Data:

- yN x k matrix of observations on k dimensional data

Prioralpha:

- Istarminexpected number of components at lower bound of support of alpha
- Istarmaxexpected number of components at upper bound of support of alpha

- powerpower parameter for alpha prior

lambda\\_hyper:

- alimdefines support of a distribution,def:c(.01,10)
- nulimdefines support of nu distribution, def:c(.01,3)
- vlimdefines support of v distribution, def:c(.1,4)

Mcmc:

- Rnumber of mcmc draws
- keepthinning parm, keep every keepth draw
- maxuniqstorage constraint on the number of unique components
- SCALEshould data be scaled by mean,std deviation before posterior draws, def: TRUE
- gridsizenumber of discrete points for hyperparameter priors,def: 20

output:

the basic output are draws from the predictive distribution of the data in the object, nmix. The average of these draws is the Bayesian analogue of a density estimate.

nmix:

- probdrawR/keep x 1 matrix of 1s
- zdrawR/keep x N matrix of draws of indicators of which component each obs is assigned to
- compdrawR/keep list of draws of normals

Output of the components is in the form of a list of lists.

compdraw[[i]] is ith draw – list of lists.

compdraw[[i]][[1]] is list of parms for a draw from predictive.

compdraw[[i]][[1]][[1]] is the mean vector. compdraw[[i]][[1]][[2]] is the inverse of Cholesky root.  $\Sigma = t(R) \% \% R$ ,  $R^{-1} = \text{compdraw}[[i]][[1]][[2]]$ .

## Value

nmix	a list containing: probdraw,zdraw,compdraw
alphadraw	vector of draws of DP process tightness parameter
nudraw	vector of draws of base prior hyperparameter
adraw	vector of draws of base prior hyperparameter
vdraw	vector of draws of base prior hyperparameter

## Note

we parameterize the prior on  $\Sigma_i$  such that  $\text{mode}(\Sigma_i) = nu/(nu + 2)vI$ . The support of nu enforces valid IW density;  $nulim[1] > 0$

We use the structure for nmix that is compatible with the bayesm routines for finite mixtures of normals. This allows us to use the same summary and plotting methods.

The default choices of alim,nulim, and vlim determine the location and approximate size of candidate "atoms" or possible normal components. The defaults are sensible given that we scale the data. Without scaling, you want to insure that alim is set for a wide enough range of values (remember a

is a precision parameter) and the  $v$  is big enough to propose Sigma matrices wide enough to cover the data range.

A careful analyst should look at the posterior distribution of  $a$ ,  $\nu$ ,  $v$  to make sure that the support is set correctly in `alim`, `nulim`, `vlim`. In other words, if we see the posterior bunched up at one end of these support ranges, we should widen the range and rerun.

If you want to force the procedure to use many small atoms, then set `nulim` to consider only large values and set `vlim` to consider only small scaling constants. Set `Istarmax` to a large number. This will create a very "lumpy" density estimate somewhat like the classical Kernel density estimates. Of course, this is not advised if you have a prior belief that densities are relatively smooth.

### Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

### See Also

[rnmixGibbs](#), [rmixture](#), [rmixGibbs](#), [eMixMargDen](#), [momMix](#), [mixDen](#), [mixDenBi](#)

### Examples

```
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

## simulate univariate data from Chi-Sq

set.seed(66)
N=200
chisqdf=8; y1=as.matrix(rchisq(N,df=chisqdf))

## set arguments for rDPGibbs

Data1=list(y=y1)
Prioralpha=list(Istarmin=1,Istarmax=10,power=.8)
Prior1=list(Prioralpha=Prioralpha)

Mcmc=list(R=R,keep=1,maxuniq=200)

out1=rDPGibbs(Prior=Prior1,Data=Data1,Mcmc)

if(0){
## plotting examples
rgi=c(0,20); grid=matrix(seq(from=rgi[1],to=rgi[2],length.out=50),ncol=1)
deltax=(rgi[2]-rgi[1])/nrow(grid)
plot(out1$nmix,Grid=grid,Data=y1)
## plot true density with histogram
plot(range(grid[,1]),1.5*range(dchisq(grid[,1],df=chisqdf)),type="n",xlab=paste("Chisq ; ",N," obs",sep=""))
hist(y1,xlim=rgi,freq=FALSE,col="yellow",breaks=20,add=TRUE)
lines(grid[,1],dchisq(grid[,1],df=chisqdf)/(sum(dchisq(grid[,1],df=chisqdf))*deltax),col="blue",lwd=2)
}

## simulate bivariate data from the "Banana" distribution (Meng and Barnard)
banana=function(A,B,C1,C2,N,keep=10,init=10)
{ R=init*keep+N*keep
  x1=x2=0
  bimat=matrix(double(2*N),ncol=2)
```



```

    for (r in 1:R)
    { x1=rnorm(1,mean=(B*x2+C1)/(A*(x2^2)+1),sd=sqrt(1/(A*(x2^2)+1)))
      x2=rnorm(1,mean=(B*x2+C2)/(A*(x1^2)+1),sd=sqrt(1/(A*(x1^2)+1)))
      if (r>init*keep && r%%keep==0) {mkeep=r/keep; bimat[mkeep-init,]=c(x1,x2)} }
    return(bimat)
  }

  set.seed(66)
  nvar2=2
  A=0.5; B=0; C1=C2=3
  y2=banana(A=A,B=B,C1=C1,C2=C2,1000)

  Data2=list(y=y2)
  Prioralpha=list(Istarmin=1,Istarmax=10,power=.8)
  Prior2=list(Prioralpha=Prioralpha)
  Mcmc=list(R=R,keep=1,maxuniq=200)

  out2=rDPGibbs(Prior=Prior2,Data=Data2,Mcmc)

  if(0){
    ## plotting examples

    rx1=range(y2[,1]); rx2=range(y2[,2])
    x1=seq(from=rx1[1],to=rx1[2],length.out=50)
    x2=seq(from=rx2[1],to=rx2[2],length.out=50)
    grid=cbind(x1,x2)

    plot(out2$nmix,Grid=grid,Data=y2)

    ## plot true bivariate density
    tden=matrix(double(50*50),ncol=50)
    for (i in 1:50){ for (j in 1:50)
      {tden[i,j]=exp(-0.5*(A*(x1[i]^2)*(x2[j]^2)+(x1[i]^2)+(x2[j]^2)-2*B*x1[i]*x2[j]-2*C1*x1[i]-2*C2*x2[j]))}
    }
    tden=tden/sum(tden)
    image(x1,x2,tden,col=terrain.colors(100),xlab="",ylab="")
    contour(x1,x2,tden,add=TRUE,drawlabels=FALSE)
    title("True Density")
  }

```

## Description

rhierBinLogit implements an MCMC algorithm for hierarchical binary logits with a normal heterogeneity distribution. This is a hybrid sampler with a RW Metropolis step for unit-level logit parameters.

rhierBinLogit is designed for use on choice-based conjoint data with partial profiles. The Design matrix is based on differences of characteristics between two alternatives. See Appendix A of *Bayesian Statistics and Marketing* for details.

**Usage**

```
rhierBinLogit(Data, Prior, Mcmc)
```

**Arguments**

Data	list(lgtdata,Z) (note: Z is optional)
Prior	list(Deltabar,ADelta,nu,V) (note: all are optional)
Mcmc	list(sbeta,R,keep) (note: all but R are optional)

**Details**

Model:

$y_{hi} = 1$  with  $pr = \exp(x'_{hi} \beta_{hi}) / (1 + \exp(x'_{hi} \beta_{hi}))$ .  $\beta_{hi}$  is  $nvar \times 1$ .  
 $h=1, \dots, \text{length}(\text{lgtdata})$  units or "respondents" for survey data.

$\beta_{hi} = Z\Delta[h,] + u_h$ .

Note: here ZDelta refers to  $Z\%*\%Delta$ , ZDelta[h,] is hth row of this product.

Delta is an  $nz \times nvar$  array.

$u_h \sim N(0, V_{\beta_{hi}})$ .

Priors:

$\Delta = \text{vec}(\Delta) \sim N(\text{vec}(\text{Deltabar}), V_{\beta_{hi}}(x) A\Delta^{-1})$

$V_{\beta_{hi}} \sim IW(nu, V)$

Lists contain:

- lgtdatalist of lists with each cross-section unit MNL data
- lgtdata[[h]]\$y  $n_h$  vector of binary outcomes (0,1)
- lgtdata[[h]]\$X  $n_h$  by  $nvar$  design matrix for hth unit
- Deltabarnz  $x \times nvar$  matrix of prior means (def: 0)
- ADelta prior prec matrix (def: .01I)
- nu d.f. parm for IW prior on norm comp Sigma (def:  $nvar+3$ )
- V pds location parm for IW prior on norm comp Sigma (def: nuI)
- sbeta scaling parm for RW Metropolis (def: .2)
- R number of MCMC draws
- keep MCMC thinning parm: keep every keepth draw (def: 1)

**Value**

a list containing:

Deltadraw	R/keep $x \times nz \times nvar$ matrix of draws of Delta
betadraw	$nlt \times nvar \times R/keep$ array of draws of betas
Vbetadraw	R/keep $x \times nvar \times nvar$ matrix of draws of Vbeta
llike	R/keep vector of log-likelihood values
reject	R/keep vector of reject rates over $nlt$ units

**Note**

Some experimentation with the Metropolis scaling paramter (sbeta) may be required.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 5.

<http://www.perossi.org/home/bsm-1>

**See Also**

[rhierMnlRwMixture](#)

**Examples**

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=10000} else {R=10}

set.seed(66)
nvar=5                      ## number of coefficients
nlgt=1000                   ## number of cross-sectional units
nobs=10                     ## number of observations per unit
nz=2                        ## number of regressors in mixing distribution

## set hyper-parameters
##      B=ZDelta + U

Z=matrix(c(rep(1,nlgt),runif(nlgt,min=-1,max=1)),nrow=nlgt,ncol=nz)
Delta=matrix(c(-2,-1,0,1,2,-1,1,-.5,.5,0),nrow=nz,ncol=nvar)
iota=matrix(1,nrow=nvar,ncol=1)
Vbeta=diag(nvar)+.5*iota%*%t(iota)

## simulate data
lgtdata=NULL

for (i in 1:nlgt)
{ beta=t(Delta)%*%Z[i,]+as.vector(t(chol(Vbeta))%*%rnorm(nvar))
  X=matrix(runif(nobs*nvar),nrow=nobs,ncol=nvar)
  prob=exp(X%*%beta)/(1+exp(X%*%beta))
  unif=runif(nobs,0,1)
  y=ifelse(unif<prob,1,0)
  lgtdata[[i]]=list(y=y,X=X,beta=beta)
}

out=rhierBinLogit(Data=list(lgtdata=lgtdata,Z=Z),Mcmc=list(R=R))

cat("Summary of Delta draws",fill=TRUE)
summary(out$Deltadraw,tvalues=as.vector(Delta))
cat("Summary of Vbeta draws",fill=TRUE)
summary(out$Vbetadraw,tvalues=as.vector(Vbeta[upper.tri(Vbeta,diag=TRUE)]))

if(0){
## plotting examples
plot(out$Deltadraw,tvalues=as.vector(Delta))
plot(out$betadraw)
plot(out$Vbetadraw,tvalues=as.vector(Vbeta[upper.tri(Vbeta,diag=TRUE)]))
}
```

}

---

rhierLinearMixture      *Gibbs Sampler for Hierarchical Linear Model*


---

## Description

rhierLinearMixture implements a Gibbs Sampler for hierarchical linear models with a mixture of normals prior.

## Usage

```
rhierLinearMixture(Data, Prior, Mcmc)
```

## Arguments

Data	list(regdata,Z) (Z optional).
Prior	list(deltabar,Ad,mubar,Amu,nu,V,nu.e,ssq,ncomp) (all but ncomp are optional).
Mcmc	list(R,keep) (R required).

## Details

Model: length(regdata) regression equations.

$y_i = X_i \beta_{i\cdot} + e_i$ .  $e_i \sim N(0, \tau_{i\cdot})$ . nvar X vars in each equation.

Priors:

$\tau_{i\cdot} \sim \text{nu.e} * \text{ssq}_i / \chi^2_{\text{nu.e}}$ .  $\tau_{i\cdot}$  is the variance of  $e_i$ .

$\beta_{i\cdot} = \text{ZDelta}[i,] + u_i$ .

Note: here ZDelta refers to  $\text{Z} \% \% \text{D}$ , ZDelta[i,] is ith row of this product.

Delta is an nz x nvar array.

$u_i \sim N(\mu_{i\cdot}, \text{Sigma}_{i\cdot})$ .  $i\cdot \sim \text{multinomial}(pvec)$ .

$pvec \sim \text{dirichlet}(a)$

$\text{delta} = \text{vec}(\text{Delta}) \sim N(\text{deltabar}, A_d^{-1})$

$\mu_{j\cdot} \sim N(\text{mubar}, \text{Sigma}_j(x) A_{\mu}^{-1})$

$\text{Sigma}_j \sim \text{IW}(\text{nu}, V)$

List arguments contain:

- regdata list of lists with X,y matrices for each of length(regdata) regressions
- regdata[[i]]\$X X matrix for equation i
- regdata[[i]]\$y y vector for equation i
- deltabarnz\*nvar vector of prior means (def: 0)
- Ad prior prec matrix for vec(Delta) (def: .01I)
- mubar nvar x 1 prior mean vector for normal comp mean (def: 0)
- Amu prior precision for normal comp mean (def: .01I)

- nu d.f. parm for IW prior on norm comp Sigma (def: nvar+3)
- V pds location parm for IW prior on norm comp Sigma (def: nuI)
- nu.e d.f. parm for regression error variance prior (def: 3)
- ssq scale parm for regression error var prior (def: var( $y_i$ ))
- ncomp number of components used in normal mixture
- R number of MCMC draws
- keep MCMC thinning parm: keep every keepth draw (def: 1)

### Value

a list containing

taudraw	R/keep x nreg array of error variance draws
betadraw	nreg x nvar x R/keep array of individual regression coef draws
Deltadraw	R/keep x nz x nvar array of Deltadraws
nmix	list of three elements, (probdraw, NULL, compdraw)

### Note

More on probdraw component of nmix return value list:  
this is an R/keep by ncomp array of draws of mixture component probs (pvec)  
More on compdraw component of nmix return value list:

**compdraw[[i ]]** the ith draw of components for mixtures

**compdraw[[i ][[j]]]** ith draw of the jth normal mixture comp

**compdraw[[i ][[j]][[1]]]** ith draw of jth normal mixture comp mean vector

**compdraw[[i ][[j]][[2]]]** ith draw of jth normal mixture cov parm (rooti)

Note: Z should **not** include an intercept and should be centered for ease of interpretation.

Be careful in assessing the prior parameter, Amu. .01 can be too small for some applications. See Rossi et al, chapter 5 for full discussion.

### Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

### References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 3.  
<http://www.perossi.org/home/bsm-1>

### See Also

[rhierLinearModel](#)

**Examples**

```

##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

set.seed(66)
nreg=300; nobs=500; nvar=3; nz=2

Z=matrix(runif(nreg*nz),ncol=nz)
Z=t(t(Z)-apply(Z,2,mean))
Delta=matrix(c(1,-1,2,0,1,0),ncol=nz)
tau0=.1
iota=c(rep(1,nobs))

## create arguments for rmixture

tcomps=NULL
a=matrix(c(1,0,0,0.5773503,1.1547005,0,-0.4082483,0.4082483,1.2247449),ncol=3)
tcomps[[1]]=list(mu=c(0,-1,-2),rooti=a)
tcomps[[2]]=list(mu=c(0,-1,-2)*2,rooti=a)
tcomps[[3]]=list(mu=c(0,-1,-2)*4,rooti=a)
tpvec=c(.4,.2,.4)

regdata=NULL # simulated data with Z
betas=matrix(double(nreg*nvar),ncol=nvar)
tind=double(nreg)

for (reg in 1:nreg) {
  tempout=rmixture(1,tpvec,tcomps)
  betas[reg,]=Delta%*%Z[reg,]+as.vector(tempout$x)
  tind[reg]=tempout$z
  X=cbind(iota,matrix(runif(nobs*(nvar-1)),ncol=(nvar-1)))
  tau=tau0*runif(1,min=0.5,max=1)
  y=X%*%betas[reg,]+sqrt(tau)*rnorm(nobs)
  regdata[[reg]]=list(y=y,X=X,beta=betas[reg,],tau=tau)
}

## run rhierLinearMixture

Data1=list(regdata=regdata,Z=Z)
Prior1=list(ncomp=3)
Mcmc1=list(R=R,keep=1)

out1=rhierLinearMixture(Data=Data1,Prior=Prior1,Mcmc=Mcmc1)

cat("Summary of Delta draws",fill=TRUE)
summary(out1$Deltadraw,tvalues=as.vector(Delta))
cat("Summary of Normal Mixture Distribution",fill=TRUE)
summary(out1$nmix)

if(0){
## plotting examples
plot(out1$betadraw)
plot(out1$nmix)
plot(out1$Deltadraw)
}

```

rhierLinearModel

*Gibbs Sampler for Hierarchical Linear Model***Description**

rhierLinearModel implements a Gibbs Sampler for hierarchical linear models with a normal prior.

**Usage**

```
rhierLinearModel(Data, Prior, Mcmc)
```

**Arguments**

Data	list(regdata,Z) (Z optional).
Prior	list(Deltabar,A,nu.e,ssq,nu,V) (optional).
Mcmc	list(R,keep) (R required).

**Details**

Model:  $\text{length}(\text{regdata})$  regression equations.  
 $y_i = X_i \beta_i + e_i$ .  $e_i \sim N(0, \tau_{u_i})$ . nvar X vars in each equation.

Priors:

$\tau_{u_i} \sim \text{nu.e} * \text{ssq}_i / \chi^2_{\text{nu.e}}$ .  $\tau_{u_i}$  is the variance of  $e_i$ .

$\beta_i \sim N(\text{ZDelta}[i,], V_{\beta_i})$ .

Note: ZDelta is the matrix  $Z * \text{Delta}$ ; [i,] refers to ith row of this product.

$\text{vec}(\text{Delta})$  given  $V_{\beta_i} \sim N(\text{vec}(\text{Deltabar}), V_{\beta_i}(x)A^{-1})$ .

$V_{\beta_i} \sim IW(\text{nu}, V)$ .

$\text{Delta}, \text{Deltabar}$  are  $\text{nz} \times \text{nvar}$ .  $A$  is  $\text{nz} \times \text{nz}$ .  $V_{\beta_i}$  is  $\text{nvar} \times \text{nvar}$ .

Note: if you don't have any z vars, set  $Z=\text{iota}(\text{nreg} \times 1)$ .

List arguments contain:

- regdata list of lists with X,y matrices for each of  $\text{length}(\text{regdata})$  regressions
- regdata[[i]]\$X X matrix for equation i
- regdata[[i]]\$y y vector for equation i
- Deltabar  $\text{nz} \times \text{nvar}$  matrix of prior means (def: 0)
- A  $\text{nz} \times \text{nz}$  matrix for prior precision (def: .01I)
- nu.e d.f. parm for regression error variance prior (def: 3)
- ssq scale parm for regression error var prior (def:  $\text{var}(y_i)$ )
- nu d.f. parm for Vbeta prior (def:  $\text{nvar}+3$ )
- V Scale location matrix for Vbeta prior (def:  $\text{nu} * I$ )
- R number of MCMC draws
- keep MCMC thinning parm: keep every keepth draw (def: 1)

**Value**

a list containing

betadraw	nreg x nvar x R/keep array of individual regression coef draws
taudraw	R/keep x nreg array of error variance draws
Deltadraw	R/keep x nz x nvar array of Deltadraws
Vbetadraw	R/keep x nvar*nvar array of Vbeta draws

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 3.

<http://www.perossi.org/home/bsm-1>

**See Also**

[rhierLinearMixture](#)

**Examples**

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

nreg=100; nobs=100; nvar=3
Vbeta=matrix(c(1,.5,0,.5,2,.7,0,.7,1),ncol=3)
Z=cbind(c(rep(1,nreg)),3*runif(nreg)); Z[,2]=Z[,2]-mean(Z[,2])
nz=ncol(Z)
Delta=matrix(c(1,-1,2,0,1,0),ncol=2)
Delta=t(Delta) # first row of Delta is means of betas
Beta=matrix(rnorm(nreg*nvar),nrow=nreg)%*%chol(Vbeta)+Z%*%Delta
tau=.1
iota=c(rep(1,nobs))
regdata=NULL
for (reg in 1:nreg) { X=cbind(iota,matrix(runif(nobs*(nvar-1)),ncol=(nvar-1)))
y=X%*%Beta[reg,]+sqrt(tau)*rnorm(nobs); regdata[[reg]]=list(y=y,X=X) }

Data1=list(regdata=regdata,Z=Z)
Mcmc1=list(R=R,keep=1)
out=rhierLinearModel(Data=Data1,Mcmc=Mcmc1)

cat("Summary of Delta draws",fill=TRUE)
summary(out$Deltadraw,tvalues=as.vector(Delta))
cat("Summary of Vbeta draws",fill=TRUE)
summary(out$Vbetadraw,tvalues=as.vector(Vbeta[upper.tri(Vbeta,diag=TRUE)]))

if(0){
## plotting examples
plot(out$betadraw)
plot(out$Deltadraw)
}
```



rhierMnlDP

*MCMC Algorithm for Hierarchical Multinomial Logit with Dirichlet Process Prior Heterogeneity*

## Description

rhierMnlDP is a MCMC algorithm for a hierarchical multinomial logit with a Dirichlet Process Prior for the distribution of heterogeneity. A base normal model is used so that the DP can be interpreted as allowing for a mixture of normals with as many components as there are panel units. This is a hybrid Gibbs Sampler with a RW Metropolis step for the MNL coefficients for each panel unit. This procedure can be interpreted as a Bayesian semi-parametric method in the sense that the DP prior can accommodate heterogeneity of an unknown form.

## Usage

```
rhierMnlDP(Data, Prior, Mcmc)
```

## Arguments

Data	list(p,lgtdata,Z) ( Z is optional)
Prior	list(deltabar,Ad,Prioralpha,lambda\_hyper) (all are optional)
Mcmc	list(s,w,R,keep) (R required)

## Details

Model:

$$y_i \sim MNL(X_i, \beta_{\theta_i}). \quad i=1, \dots, \text{length}(\text{lgtdata}). \quad \theta_{\theta_i} \text{ is } nvar \times 1.$$

$$\beta_{\theta_i} = Z\Delta[i,] + u_i.$$

Note: here ZDelta refers to  $Z\% \% D$ , ZDelta[i,] is ith row of this product.

Delta is an  $nz \times nvar$  array.

$$\beta_{\theta_i} \sim N(\mu_i, \Sigma_i).$$

Priors:

$$\theta_{\theta_i} = (\mu_i, \Sigma_i) \sim DP(G_0(\lambda), \alpha)$$

$$G_0(\lambda) :$$

$$\mu_i | \Sigma_i \sim N(0, \Sigma_i(x)a^{-1})$$

$$\Sigma_i \sim IW(nu, nu * v * I)$$

$$\lambda(a, nu, v) :$$

$$a \sim \text{uniform}[\text{alim}[1], \text{alim}[2]]$$

$$nu \sim \text{dim}(\text{data}) - 1 + \exp(z)$$

$$z \sim \text{uniform}[\text{dim}(\text{data}) - 1 + \text{nulim}[1], \text{nulim}[2]]$$

$$v \sim \text{uniform}[\text{vlim}[1], \text{vlim}[2]]$$

$$\alpha \sim (1 - (\alpha - \alpha_{\min}) / (\alpha_{\max} - \alpha_{\min}))^{\text{power}}$$

$\alpha_{\min}$  = alphamin then expected number of components = Istarmin

$\alpha_{\max}$  = alphamax then expected number of components = Istarmax

Lists contain:

Data:

- `p` is number of choice alternatives
- `lgtdata` list of lists with each cross-section unit MNL data
- `lgtdata[[i]]$y`  $n_i$  vector of multinomial outcomes (1,...,m)
- `lgtdata[[i]]$X`  $n_i$  by `nvar` design matrix for `i`th unit

Prior:

- `deltabarnz*nvar` vector of prior means (def: 0)
- `Ad` prior prec matrix for `vec(D)` (def: .01I)

Prioralpha:

- `Istarm` expected number of components at lower bound of support of `alpha` (def: 1)
- `Istarmax` expected number of components at upper bound of support of `alpha` (def: `min(50,.1*nlgt)`)
- `powerpower` parameter for `alpha` prior (def: .8)

`lambda\_hyper`:

- `alim` defines support of `a` distribution, def: c(.01,2)
- `nulim` defines support of `nu` distribution, def: c(.01,3)
- `vlim` defines support of `v` distribution, def: c(.1,4)

Mcmc:

- `Rnumber` of mcmc draws
- `keepthinning` parm, keep every `keepth` draw
- `maxuniqstorage` constraint on the number of unique components
- `gridsize` number of discrete points for hyperparameter priors, def: 20

## Value

a list containing:

<code>Deltadraw</code>	R/keep <code>x nz*nvar</code> matrix of draws of <code>Delta</code> , first row is initial value
<code>betadraw</code>	<code>nlgt x nvar x R</code> /keep array of draws of <code>betas</code>
<code>nmix</code>	list of 3 components, <code>probdraw</code> , <code>NULL</code> , <code>compdraw</code>
<code>adraw</code>	R/keep draws of hyperparm <code>a</code>
<code>vdraw</code>	R/keep draws of hyperparm <code>v</code>
<code>nudraw</code>	R/keep draws of hyperparm <code>nu</code>
<code>Istardraw</code>	R/keep draws of number of unique components
<code>alphadraw</code>	R/keep draws of number of DP tightness parameter
<code>loglike</code>	R/keep draws of log-likelihood

**Note**

As is well known, Bayesian density estimation involves computing the predictive distribution of a "new" unit parameter,  $\theta_{n+1}$  (here "n"=nlgt). This is done by averaging the normal base distribution over draws from the distribution of  $\theta_{n+1}$  given  $\theta_1, \dots, \theta_n, \alpha, \lambda, \text{Data}$ . To facilitate this, we store those draws from the predictive distribution of  $\theta_{n+1}$  in a list structure compatible with other bayesm routines that implement a finite mixture of normals.

More on nmix list:

contains the draws from the predictive distribution of a "new" observations parameters. These are simply the parameters of one normal distribution. We enforce compatibility with a mixture of k components in order to utilize generic summary plotting functions.

Therefore, probdraw is a vector of ones. zdraw (indicator draws) is omitted as it is not necessary for density estimation. compdraw contains the draws of the  $\theta_{n+1}$  as a list of list of lists.

More on compdraw component of return value list:

- compdraw[[i]]ith draw of components for mixtures
- compdraw[[i]][[1]]ith draw of the thetanp1
- compdraw[[i]][[1]][[1]]ith draw of mean vector
- compdraw[[i]][[1]][[2]]ith draw of parm (rooti)

We parameterize the prior on  $\Sigma_i$  such that  $\text{mode}(\Sigma) = \nu/(\nu + 2)vI$ . The support of  $\nu$  enforces a non-degenerate IW density;  $\text{nulim}[1] > 0$ .

The default choices of alim, nulim, and vlim determine the location and approximate size of candidate "atoms" or possible normal components. The defaults are sensible given a reasonable scaling of the X variables. You want to insure that alim is set for a wide enough range of values (remember a is a precision parameter) and the v is big enough to propose Sigma matrices wide enough to cover the data range.

A careful analyst should look at the posterior distribution of a, nu, v to make sure that the support is set correctly in alim, nulim, vlim. In other words, if we see the posterior bunched up at one end of these support ranges, we should widen the range and rerun.

If you want to force the procedure to use many small atoms, then set nulim to consider only large values and set vlim to consider only small scaling constants. Set alphamax to a large number. This will create a very "lumpy" density estimate somewhat like the classical Kernel density estimates. Of course, this is not advised if you have a prior belief that densities are relatively smooth.

Note: Z should **not** include an intercept and is centered for ease of interpretation.

Large R values may be required (>20,000).

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 5.  
<http://www.perossi.org/home/bsm-1>

**See Also**

[rhierMnlRwMixture](#)

## Examples

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=20000} else {R=10}

set.seed(66)
p=3                                # num of choice alterns
ncoef=3
nlgt=300                           # num of cross sectional units
nz=2
Z=matrix(runif(nz*nlgt),ncol=nz)
Z=t(t(Z)-apply(Z,2,mean))          # demean Z
ncomp=3                            # no of mixture components
Delta=matrix(c(1,0,1,0,1,2),ncol=2)
comps=NULL
comps[[1]]=list(mu=c(0,-1,-2),rooti=diag(rep(2,3)))
comps[[2]]=list(mu=c(0,-1,-2)*2,rooti=diag(rep(2,3)))
comps[[3]]=list(mu=c(0,-1,-2)*4,rooti=diag(rep(2,3)))
pvec=c(.4,.2,.4)

simnmlwX= function(n,X,beta) {
  ## simulate from MNL model conditional on X matrix
  k=length(beta)
  Xbeta=X%%beta
  j=nrow(Xbeta)/n
  Xbeta=matrix(Xbeta,byrow=TRUE,ncol=j)
  Prob=exp(Xbeta)
  iota=c(rep(1,j))
  denom=Prob%%iota
  Prob=Prob/as.vector(denom)
  y=vector("double",n)
  ind=1:j
  for (i in 1:n)
    {yvec=rmultinom(1,1,Prob[i,]); y[i]=ind%%yvec}
  return(list(y=y,X=X,beta=beta,prob=Prob))
}

## simulate data with a mixture of 3 normals
simlgtdata=NULL
ni=rep(50,300)
for (i in 1:nlgt)
{
  betai=Delta%%Z[i,]+as.vector(rmixture(1,pvec,comps)$x)
  Xa=matrix(runif(ni[i]*p,min=-1.5,max=0),ncol=p)
  X=createX(p,na=1,nd=NULL,Xa=Xa,Xd=NULL,base=1)
  outa=simnmlwX(ni[i],X,betai)
  simlgtdata[[i]]=list(y=outa$y,X=X,beta=betai)
}

## plot betas
if(1){
  ## set if(1) above to produce plots
  bmat=matrix(0,nlgt,ncoef)
  for(i in 1:nlgt) {bmat[i,]=simlgtdata[[i]]$beta}
  par(mfrow=c(ncoef,1))
  for(i in 1:ncoef) hist(bmat[,i],breaks=30,col="magenta")
}
```

```

## set Data and Mcmc lists
keep=5
Mcmc1=list(R=R,keep=keep)
Data1=list(p=p,lgtdata=simlgtdata,Z=Z)

out=rhierMnlDP(Data=Data1,Mcmc=Mcmc1)

cat("Summary of Delta draws",fill=TRUE)
summary(out$Deltadraw,tvalues=as.vector(Delta))

if(0) {
## plotting examples
plot(out$betadraw)
plot(out$nmix)
}

```

---

rhierMnlRwMixture	<i>MCMC Algorithm for Hierarchical Multinomial Logit with Mixture of Normals Heterogeneity</i>
-------------------	--

---

## Description

rhierMnlRwMixture is a MCMC algorithm for a hierarchical multinomial logit with a mixture of normals heterogeneity distribution. This is a hybrid Gibbs Sampler with a RW Metropolis step for the MNL coefficients for each panel unit.

## Usage

```
rhierMnlRwMixture(Data, Prior, Mcmc)
```

## Arguments

Data	list(p,lgtdata,Z) ( Z is optional)
Prior	list(a,deltabar,Ad,mubar,Amu,nu,V,ncomp) (all but ncomp are optional)
Mcmc	list(s,w,R,keep) (R required)

## Details

Model:

$y_i \sim MNL(X_i, \beta_{\theta_i})$ .  $i=1, \dots, \text{length}(\text{lgtdata})$ .  $\theta_{\theta_i}$  is  $nvar \times 1$ .

$\beta_{\theta_i} = Z\Delta[i,] + u_i$ .

Note: here ZDelta refers to  $Z\%*\%D$ , ZDelta[i,] is ith row of this product.

Delta is an  $nz \times nvar$  array.

$u_i \sim N(\mu_{ind}, \Sigma_{ind})$ .  $ind \sim \text{multinomial}(pvec)$ .

Priors:

$pvec \sim \text{dirichlet}(a)$

$\Delta = \text{vec}(\Delta) \sim N(\text{deltabar}, A_d^{-1})$

$\mu_j \sim N(\text{mubar}, \Sigma_j(x) A_{\mu}^{-1})$

$$\text{Sigma}_j \sim \text{IW}(\text{nu}, \text{V})$$

Lists contain:

- p p is number of choice alternatives
- lgtdatalist of lists with each cross-section unit MNL data
- lgtdata[[i]]\$y  $n_i$  vector of multinomial outcomes (1, . . . , m)
- lgtdata[[i]]\$X  $n_i \times p$  by nvar design matrix for ith unit
- avector of length ncomp of Dirichlet prior parms (def: rep(5, ncomp))
- deltabarnz\*nvar vector of prior means (def: 0)
- Ad prior prec matrix for vec(D) (def: .01I)
- mubar nvar x 1 prior mean vector for normal comp mean (def: 0)
- Amu prior precision for normal comp mean (def: .01I)
- nu d.f. parm for IW prior on norm comp Sigma (def: nvar+3)
- V pds location parm for IW prior on norm comp Sigma (def: nuI)
- ncomp number of components used in normal mixture
- s scaling parm for RW Metropolis (def: 2.93/sqrt(nvar))
- w fractional likelihood weighting parm (def: .1)
- R number of MCMC draws
- keep MCMC thinning parm: keep every keepth draw (def: 1)

### Value

a list containing:

Deltadraw	R/keep x nz*nvar matrix of draws of Delta, first row is initial value
betadraw	nlgt x nvar x R/keep array of draws of betas
nmix	list of 3 components, probdraw, NULL, compdraw
loglike	log-likelihood for each kept draw (length R/keep)

### Note

More on probdraw component of nmix list:

R/keep x ncomp matrix of draws of probs of mixture components (pvec)

More on compdraw component of return value list:

- compdraw[[i]] the ith draw of components for mixtures
- compdraw[[i]][[j]] ith draw of the jth normal mixture comp
- compdraw[[i]][[j]][[1]] ith draw of jth normal mixture comp mean vector
- compdraw[[i]][[j]][[2]] ith draw of jth normal mixture cov parm (rooti)

Note: Z should **not** include an intercept and is centered for ease of interpretation.

Be careful in assessing prior parameter, Amu. .01 is too small for many applications. See Rossi et al, chapter 5 for full discussion.

Note: as of version 2.0-2 of bayesm, the fractional weight parameter has been changed to a weight between 0 and 1.  $w$  is the fractional weight on the normalized pooled likelihood. This differs from what is in Rossi et al chapter 5, i.e.

$$like_i(1 - w) + w \cdot like_{pooled}(n_i/N * w)$$

Large R values may be required (>20,000).

### Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

### References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 5.

<http://www.perossi.org/home/bsm-1>

### See Also

[rmnlIndepMetrop](#)

### Examples

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=10000} else {R=10}

set.seed(66)
p=3                                # num of choice alterns
ncoef=3
nlgt=300                          # num of cross sectional units
nz=2
Z=matrix(runif(nz*nlgt),ncol=nz)
Z=t(t(Z)-apply(Z,2,mean))         # demean Z
ncomp=3                            # no of mixture components
Delta=matrix(c(1,0,1,0,1,2),ncol=2)
comps=NULL
comps[[1]]=list(mu=c(0,-1,-2),rooti=diag(rep(1,3)))
comps[[2]]=list(mu=c(0,-1,-2)*2,rooti=diag(rep(1,3)))
comps[[3]]=list(mu=c(0,-1,-2)*4,rooti=diag(rep(1,3)))
pvec=c(.4,.2,.4)

simnmlwX= function(n,X,beta) {
  ## simulate from MNL model conditional on X matrix
  k=length(beta)
  Xbeta=X%%beta
  j=nrow(Xbeta)/n
  Xbeta=matrix(Xbeta,byrow=TRUE,ncol=j)
  Prob=exp(Xbeta)
  iota=c(rep(1,j))
  denom=Prob%*%iota
  Prob=Prob/as.vector(denom)
  y=vector("double",n)
  ind=1:j
  for (i in 1:n)
    {yvec=rmultinom(1,1,Prob[i,]); y[i]=ind%*%yvec}
  return(list(y=y,X=X,beta=beta,prob=Prob))
}
```

```

}

## simulate data
simlgtdata=NULL
ni=rep(50,300)
for (i in 1:nlgt)
{
  betai=Delta*%Z[i,]+as.vector(rmixture(1,pvec,comps)$x)
  Xa=matrix(runif(ni[i]*p,min=-1.5,max=0),ncol=p)
  X=createX(p,na=1,nd=NULL,Xa=Xa,Xd=NULL,base=1)
  outa=simmnlwX(ni[i],X,betai)
  simlgtdata[[i]]=list(y=outa$y,X=X,beta=betai)
}

## plot betas
if(0){
  ## set if(1) above to produce plots
  bmat=matrix(0,nlgt,ncoef)
  for(i in 1:nlgt) {bmat[i,]=simlgtdata[[i]]$beta}
  par(mfrow=c(ncoef,1))
  for(i in 1:ncoef) hist(bmat[,i],breaks=30,col="magenta")
}

## set parms for priors and Z
Prior1=list(ncomp=5)

keep=5
Mcmc1=list(R=R,keep=keep)
Data1=list(p=p,lgtdata=simlgtdata,Z=Z)

out=rhierMnlRwMixture(Data=Data1,Prior=Prior1,Mcmc=Mcmc1)

cat("Summary of Delta draws",fill=TRUE)
summary(out$Deltadraw,tvalues=as.vector(Delta))
cat("Summary of Normal Mixture Distribution",fill=TRUE)
summary(out$nmix)

if(0) {
  ## plotting examples
  plot(out$betadraw)
  plot(out$nmix)
}

```

## Description

rhierNegbinRw implements an MCMC strategy for the hierarchical Negative Binomial (NBD) regression model. Metropolis steps for each unit level set of regression parameters are automatically tuned by optimization. Over-dispersion parameter (alpha) is common across units.

## Usage

```
rhierNegbinRw(Data, Prior, Mcmc)
```



**Arguments**

Data	list(regdata,Z)
Prior	list(Deltabar,Adelta,nu,V,a,b)
Mcmc	list(R,keep,s\_beta,s\_alpha,c,Vbeta0,Delta0)

**Details**

Model:  $y_i \sim \text{NBD}(\text{mean}=\lambda, \text{over-dispersion}=\alpha)$ .

$\lambda = \exp(X_i \beta_i)$

Prior:  $\beta_i \sim N(\Delta' z_i, V\beta)$ .

$\text{vec}(\Delta | V\beta) \sim N(\text{vec}(\Delta_{\text{bar}}), V\beta(x) \Delta)$ .

$V\beta \sim \text{IW}(\nu, V)$ .

$\alpha \sim \text{Gamma}(a, b)$ .

note: prior mean of  $\alpha = a/b$ , variance =  $a/(b^2)$

list arguments contain:

- regdata list of lists with data on each of nreg units
- regdata[[i]]\$X nob<sub>s</sub> × nvar matrix of X variables
- regdata[[i]]\$y nob<sub>s</sub> × 1 vector of count responses
- Znreg x nz mat of unit chars (def: vector of ones)
- Deltabar nz x nvar prior mean matrix (def: 0)
- Adelta nz x nz pds prior prec matrix (def: .01I)
- nu d.f. parm for IWishart (def: nvar+3)
- Vlocation matrix of IWishart prior (def: nuI)
- a Gamma prior parm (def: .5)
- b Gamma prior parm (def: .1)
- R number of MCMC draws
- keep MCMC thinning parm: keep every keepth draw (def: 1)
- s\_beta scaling for beta | alpha RW inc cov (def: 2.93/sqrt(nvar))
- s\_alpha scaling for alpha | beta RW inc cov (def: 2.93)
- c fractional likelihood weighting parm (def: 2)
- Vbeta0 starting value for Vbeta (def: I)
- Delta0 starting value for Delta (def: 0)

**Value**

a list containing:

llike	R/keep vector of values of log-likelihood
betadraw	nreg x nvar x R/keep array of beta draws
alphadraw	R/keep vector of alpha draws
acceptrbeta	acceptance rate of the beta draws
acceptralpha	acceptance rate of the alpha draws

**Note**

The NBD regression encompasses Poisson regression in the sense that as  $\alpha$  goes to infinity the NBD distribution tends to the Poisson.

For "small" values of  $\alpha$ , the dependent variable can be extremely variable so that a large number of observations may be required to obtain precise inferences.

For ease of interpretation, we recommend demeaning  $Z$  variables.

**Author(s)**

Sridhar Narayanam & Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 5.

<http://www.perossi.org/home/bsm-1>

**See Also**

[rnegbinRw](#)

**Examples**

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}
##
set.seed(66)
simnegbin =
function(X, beta, alpha) {
# Simulate from the Negative Binomial Regression
lambda = exp(X %*% beta)
y=NULL
for (j in 1:length(lambda))
  y = c(y,rnbinom(1,mu = lambda[j],size = alpha))
return(y)
}

nreg = 100      # Number of cross sectional units
T = 50         # Number of observations per unit
nobs = nreg*T
nvar=2         # Number of X variables
nz=2          # Number of Z variables

# Construct the Z matrix
Z = cbind(rep(1,nreg),rnorm(nreg,mean=1,sd=0.125))

Delta = cbind(c(4,2), c(0.1,-1))
alpha = 5
Vbeta = rbind(c(2,1),c(1,2))

# Construct the regdata (containing X)
simnegbindata = NULL
for (i in 1:nreg) {
  betai = as.vector(Z[i,]%*%Delta) + chol(Vbeta)%*%rnorm(nvar)
  X = cbind(rep(1,T),rnorm(T,mean=2,sd=0.25))
```

```

    simnegbindata[[i]] = list(y=simnegbin(X,betai,alpha), X=X,beta=betai)
  }

  Beta = NULL
  for (i in 1:nreg) {Beta=rbind(Beta,matrix(simnegbindata[[i]]$beta,nrow=1))}

  Data1 = list(regdata=simnegbindata, Z=Z)
  Mcmc1 = list(R=R)

  out = rhierNegbinRw(Data=Data1, Mcmc=Mcmc1)

  cat("Summary of Delta draws",fill=TRUE)
  summary(out$Deltadraw,tvalues=as.vector(Delta))
  cat("Summary of Vbeta draws",fill=TRUE)
  summary(out$Vbetadraw,tvalues=as.vector(Vbeta[upper.tri(Vbeta,diag=TRUE)]))
  cat("Summary of alpha draws",fill=TRUE)
  summary(out$alpha,tvalues=alpha)

  if(0){
  ## plotting examples
  plot(out$betadraw)
  plot(out$alpha,tvalues=alpha)
  plot(out$Deltadraw,tvalues=as.vector(Delta))
  }

```

rivDP

*Linear "IV" Model with DP Process Prior for Errors*

## Description

rivDP is a Gibbs Sampler for a linear structural equation with an arbitrary number of instruments. rivDP uses a mixture of normals for the structural and reduced form equation implemented with a Dirichlet Process Prior.

## Usage

```
rivDP(Data, Prior, Mcmc)
```

## Arguments

Data	list(z,w,x,y)
Prior	list(md,Ad,mbg,Abg,lambd,Prioralpha) (optional)
Mcmc	list(R,keep,SCALE) (R required)

## Details

### Model:

$$x = z' \delta + e1.$$

$$y = \beta x + w' \gamma + e2.$$

$$e1, e2 \sim N(\theta_i). \theta_i \text{ represents } \mu_i, \Sigma_i$$

Note: Error terms have non-zero means. DO NOT include intercepts in the z or w matrices. This is different from rivGibbs which requires intercepts to be included explicitly.

Priors:

$$\delta \sim N(\text{md}, \text{Ad}^{-1}). \text{vec}(\beta, \gamma) \sim N(\text{mbg}, \text{Abg}^{-1})$$

$$\theta_i \sim G$$

$$G \sim DP(\alpha, G_0)$$

$G_0$  is the natural conjugate prior for  $(\mu, \Sigma)$ :

$$\Sigma \sim IW(\nu, \nu I) \text{ and } \mu | \Sigma \sim N(0, 1/\alpha \mu \Sigma)$$

These parameters are collected together in the list `lambda`. It is highly recommended that you use the default settings for these hyper-parameters.

$$\alpha \sim (1 - (\alpha - \alpha_{\min})/(\alpha_{\max} - \alpha_{\min}))^{\text{power}}$$

where  $\alpha_{\min}$  and  $\alpha_{\max}$  are set using the arguments in the reference below. It is highly recommended that you use the default values for the hyperparameters of the prior on  $\alpha$

List arguments contain:

- `z` matrix of obs on instruments
- `y` vector of obs on lhs var in structural equation
- `x` "endogenous" var in structural eqn
- `w` matrix of obs on "exogenous" vars in the structural eqn
- `md` prior mean of  $\delta$  (def: 0)
- `Ad` pds prior prec for prior on  $\delta$  (def: .01I)
- `mbg` prior mean vector for prior on  $\beta, \gamma$  (def: 0)
- `Abg` pds prior prec for prior on  $\beta, \gamma$  (def: .01I)
- `lambda` list of hyperparameters for  $\theta$  prior- use default settings
- `Prioralpha` list of hyperparameters for  $\theta$  prior- use default settings
- `R` number of MCMC draws
- `keep` MCMC thinning parm: keep every `keepth` draw (def: 1)
- `SCALE` scale data, def: TRUE
- `gridsize` `gridsize` parm for  $\alpha$  draws (def: 20)

output includes object `nmix` of class "bayesm.nmix" which contains draws of predictive distribution of errors (a Bayesian analogue of a density estimate for the error terms).

`nmix`:

- `probdraw` not used
- `zdraw` not used
- `compdraw` list `R/keep` of draws from bivariate predictive for the errors

note: in `compdraw` list, there is only one component per draw

**Value**

a list containing:

deltadraw	R/keep x dim(delta) array of delta draws
betadraw	R/keep x 1 vector of beta draws
gammadraw	R/keep x dim(gamma) array of gamma draws
Istardraw	R/keep x 1 array of draws of the number of unique normal components
alphadraw	R/keep x 1 array of draws of Dirichlet Process tightness parameter
nmix	R/keep x list of draws for predictive distribution of errors

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see "A Semi-Parametric Bayesian Approach to the Instrumental Variable Problem," by Conley, Hansen, McCulloch and Rossi, Journal of Econometrics (2008).

**See Also**

rivGibbs

**Examples**

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

##
## simulate scaled log-normal errors and run
##
set.seed(66)
k=10
delta=1.5
Sigma=matrix(c(1,.6,.6,1),ncol=2)
N=1000
tbeta=4
set.seed(66)
scalefactor=.6
root=chol(scalefactor*Sigma)
mu=c(1,1)
##
## compute interquartile ranges
##
ninterq=qnorm(.75)-qnorm(.25)
error=matrix(rnorm(100000*2),ncol=2)
error=t(t(error)+mu)
Err=t(t(exp(error))-exp(mu+.5*scalefactor*diag(Sigma)))
lnNinterq=quantile(Err[,1],prob=.75)-quantile(Err[,1],prob=.25)
##
## simulate data
##
error=matrix(rnorm(N*2),ncol=2)%*%root
```

```

error=t(t(error)+mu)
Err=t(t(exp(error))-exp(mu+.5*scalefactor*diag(Sigma)))
#
# scale appropriately
Err[,1]=Err[,1]*ninterq/lnNinterq
Err[,2]=Err[,2]*ninterq/lnNinterq
z=matrix(runif(k*N),ncol=k)
x=z%%(delta*c(rep(1,k)))+Err[,1]
y=x*tbeta+Err[,2]

# set intial values for MCMC
Data = list(); Mcmc=list()
Data$z = z; Data$x=x; Data$y=y

# start MCMC and keep results
Mcmc$maxuniq=100
Mcmc$R=R
end=Mcmc$R
begin=100

out=rivDP(Data=Data,Mcmc=Mcmc)

cat("Summary of Beta draws",fill=TRUE)
summary(out$betadraw,tvalues=tbeta)

if(0){
## plotting examples
plot(out$betadraw,tvalues=tbeta)
plot(out$nmix) ## plot "fitted" density of the errors
##
}

```

---

rivGibbs

---

*Gibbs Sampler for Linear "IV" Model*


---

## Description

rivGibbs is a Gibbs Sampler for a linear structural equation with an arbitrary number of instruments.

## Usage

```
rivGibbs(Data, Prior, Mcmc)
```

## Arguments

Data	list(z,w,x,y)
Prior	list(md,Ad,mbg,Abg,nu,V) (optional)
Mcmc	list(R,keep) (R required)

## Details

Model:

$$x = z' \delta + e1.$$

$$y = \beta * x + w' \gamma + e2.$$

$$e1, e2 \sim N(0, \Sigma).$$

Note: if intercepts are desired in either equation, include vector of ones in z or w

Priors:

$$\delta \sim N(m\delta, A\delta^{-1}). \text{vec}(\beta, \gamma) \sim N(m\beta\gamma, A\beta\gamma^{-1})$$

$$\Sigma \sim \text{IW}(\nu, V)$$

List arguments contain:

- z matrix of obs on instruments
- y vector of obs on lhs var in structural equation
- x "endogenous" var in structural eqn
- w matrix of obs on "exogenous" vars in the structural eqn
- md prior mean of delta (def: 0)
- Ad pds prior prec for prior on delta (def: .01I)
- mbg prior mean vector for prior on beta,gamma (def: 0)
- Abg pds prior prec for prior on beta,gamma (def: .01I)
- nu d.f. parm for IW prior on Sigma (def: 5)
- V pds location matrix for IW prior on Sigma (def: nuI)
- R number of MCMC draws
- keep MCMC thinning parm: keep every keepth draw (def: 1)

## Value

a list containing:

deltadraw	R/keep x dim(delta) array of delta draws
betadraw	R/keep x 1 vector of beta draws
gammadraw	R/keep x dim(gamma) array of gamma draws
Sigmadraw	R/keep x 4 array of Sigma draws

## Author(s)

Rob McCulloch and Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 5.

<http://www.perossi.org/home/bsm-1>

## Examples

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

set.seed(66)
simIV = function(delta,beta,Sigma,n,z,w,gamma) {
  eps = matrix(rnorm(2*n),ncol=2) %%% chol(Sigma)
  x = z %%% delta + eps[,1]; y = beta*x + eps[,2] + w%%gamma
  list(x=as.vector(x),y=as.vector(y)) }
n = 200 ; p=1 # number of instruments
z = cbind(rep(1,n),matrix(runif(n*p),ncol=p))
w = matrix(1,n,1)
rho=.8
Sigma = matrix(c(1,rho,rho,1),ncol=2)
delta = c(1,4); beta = .5; gamma = c(1)
simiv = simIV(delta,beta,Sigma,n,z,w,gamma)

Mcmc1=list(); Data1 = list()
Data1$z = z; Data1$w=w; Data1$x=simiv$x; Data1$y=simiv$y
Mcmc1$R = R
Mcmc1$keep=1
out=rivGibbs(Data=Data1,Mcmc=Mcmc1)

cat("Summary of Beta draws",fill=TRUE)
summary(out$betadraw,tvalues=beta)
cat("Summary of Sigma draws",fill=TRUE)
summary(out$Sigmadraw,tvalues=as.vector(Sigma[upper.tri(Sigma,diag=TRUE)]))

if(0){
## plotting examples
plot(out$betadraw)
}
```

---

rmixGibbs

---

*Gibbs Sampler for Normal Mixtures w/o Error Checking*


---

## Description

rmixGibbs makes one draw using the Gibbs Sampler for a mixture of multivariate normals.

## Usage

```
rmixGibbs(y, Bbar, A, nu, V, a, p, z, comps)
```

## Arguments

y	data array - rows are obs
Bbar	prior mean for mean vector of each norm comp
A	prior precision parameter
nu	prior d.f. parm
V	prior location matrix for covariance priro
a	Dirichlet prior parms



p	prior prob of each mixture component
z	component identities for each observation – "indicators"
comps	list of components for the normal mixture

### Details

rmixGibbs is not designed to be called directly. Instead, use rnmixGibbs wrapper function.

### Value

a list containing:

p	draw mixture probabilities
z	draw of indicators of each component
comps	new draw of normal component parameters

### Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

### Author(s)

Rob McCulloch and Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

### References

For further discussion, see *Bayesian Statistics and Marketing* by Allenby, McCulloch, and Rossi, Chapter 5.  
<http://www.perossi.org/home/bsm-1>

### See Also

[rnmixGibbs](#)

---

rmixture

*Draw from Mixture of Normals*

---

### Description

rmixture simulates iid draws from a Multivariate Mixture of Normals

### Usage

```
rmixture(n, pvec, comps)
```

### Arguments

n	number of observations
pvec	ncomp x 1 vector of prior probabilities for each mixture component
comps	list of mixture component parameters

**Details**

comps is a list of length, ncomp = length(pvec). comps[[j]][[1]] is mean vector for the jth component. comps[[j]][[2]] is the inverse of the cholesky root of Sigma for that component

**Value**

A list containing ...

x	An n x length(comps[[1]][[1]]) array of iid draws
z	A n x 1 vector of indicators of which component each draw is taken from

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**See Also**

[rnmixGibbs](#)

---

rmnlIndepMetrop

*MCMC Algorithm for Multinomial Logit Model*


---

**Description**

rmnlIndepMetrop implements Independence Metropolis for the MNL.

**Usage**

```
rmnlIndepMetrop(Data, Prior, Mcmc)
```

**Arguments**

Data	list(p,y,X)
Prior	list(A,betabar) optional
Mcmc	list(R,keep,nu)

**Details**

Model:  $y \sim \text{MNL}(X, \beta)$ .  $Pr(y = j) = \exp(x'_j \beta) / \sum_k \exp(x'_k \beta)$ .

Prior:  $\beta \sim N(\text{betabar}, A^{-1})$

list arguments contain:

- pnumber of alternatives
- y nob vector of multinomial outcomes (1, ..., p)

- Xnobs\*p x nvar matrix
- A nvar x nvar pds prior prec matrix (def: .01I)
- betabar nvar x 1 prior mean (def: 0)
- R number of MCMC draws
- keep MCMC thinning parm: keep every keepth draw (def: 1)
- nu degrees of freedom parameter for independence t density (def: 6)

### Value

a list containing:

betadraw	R/keep x nvar array of beta draws
loglike	R/keep vector of loglike values for each draw
acceptr	acceptance rate of Metropolis draws

### Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

### References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 5.

<http://www.perossi.org/home/bsm-11>

### See Also

[rhierMnlRwMixture](#)

### Examples

```
##

if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

set.seed(66)
n=200; p=3; beta=c(1,-1,1.5,.5)

simmnl= function(p,n,beta) {
  # note: create X array with 2 alt.spec vars
  k=length(beta)
  X1=matrix(runif(n*p,min=-1,max=1),ncol=p)
  X2=matrix(runif(n*p,min=-1,max=1),ncol=p)
  X=createX(p,na=2,nd=NULL,Xd=NULL,Xa=cbind(X1,X2),base=1)
  Xbeta=X%*%beta # now do probs
  p=nrow(Xbeta)/n
  Xbeta=matrix(Xbeta,byrow=TRUE,ncol=p)
  Prob=exp(Xbeta)
  iota=c(rep(1,p))
  denom=Prob%*%iota
  Prob=Prob/as.vector(denom)
  # draw y
  y=vector("double",n)
  ind=1:p
```

```

    for (i in 1:n)
      { yvec=rmultinom(1,1,Prob[i,]); y[i]=ind**%yvec }
    return(list(y=y,X=X,beta=beta,prob=Prob))
  }

simout=simmnl(p,n,beta)

Data1=list(y=simout$y,X=simout$X,p=p); Mcmc1=list(R=R,keep=1)
out=rmnlIndepMetrop(Data=Data1,Mcmc=Mcmc1)

cat("Summary of beta draws",fill=TRUE)
summary(out$betadraw,tvalues=beta)

if(0){
## plotting examples
plot(out$betadraw)
}

```

rmnpGibbs

*Gibbs Sampler for Multinomial Probit***Description**

rmnpGibbs implements the McCulloch/Rossi Gibbs Sampler for the multinomial probit model.

**Usage**

```
rmnpGibbs(Data, Prior, Mcmc)
```

**Arguments**

Data	list(p, y, X)
Prior	list(betabar,A,nu,V) (optional)
Mcmc	list(beta0,sigma0,R,keep) (R required)

**Details**

model:

$w_i = X_i\beta + e$ .  $e \sim N(0, \text{Sigma})$ . note:  $w_i, e$  are  $(p-1) \times 1$ .

$y_i = j$ , if  $w_{ij} > \max(0, w_{i,-j})$   $j=1, \dots, p-1$ .  $w_{i,-j}$  means elements of  $w_i$  other than the  $j$ th.

$y_i = p$ , if all  $w_i < 0$ .

priors:

$\beta \sim N(\text{betabar}, A^{-1})$

$\text{Sigma} \sim \text{IW}(\text{nu}, V)$

to make up X matrix use [createX](#) with DIFF=TRUE.

List arguments contain

- pnumber of choices or possible multinomial outcomes

- $y_n$  x 1 vector of multinomial outcomes
- $X_n^{*(p-1)}$  x k Design Matrix
- $\text{betabark}$  x 1 prior mean (def: 0)
- $A_k$  x k prior precision matrix (def:  $.01I$ )
- $\nu$  d.f. parm for IWishart prior (def:  $(p-1) + 3$ )
- $V$  pds location parm for IWishart prior (def:  $\nu * I$ )
- $\text{beta0}$  initial value for beta
- $\text{sigma0}$  initial value for sigma
- R number of MCMC draws
- keep thinning parameter - keep every keepth draw (def: 1)

### Value

a list containing:

<code>betadraw</code>	R/keep x k array of betadraws
<code>sigmadraw</code>	R/keep x $(p-1)*(p-1)$ array of sigma draws – each row is in vector form

### Note

beta is not identified.  $\text{beta}/\sqrt{\text{sigma}_{11}}$  and  $\text{Sigma}/\text{sigma}_{11}$  are. See Allenby et al or example below for details.

### Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

### References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 4.

<http://www.perossi.org/home/bsm-11>

### See Also

[rmvpGibbs](#)

### Examples

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

set.seed(66)
p=3
n=500
beta=c(-1,1,1,2)
Sigma=matrix(c(1,.5,.5,1),ncol=2)
k=length(beta)
X1=matrix(runif(n*p,min=0,max=2),ncol=p); X2=matrix(runif(n*p,min=0,max=2),ncol=p)
X=createX(p,na=2,nd=NULL,Xa=cbind(X1,X2),Xd=NULL,DIFF=TRUE,base=p)

simnp= function(X,p,n,beta,sigma) {
```

```

indmax=function(x) {which(max(x)==x)}
Xbeta=X%*%beta
w=as.vector(crossprod(chol(sigma),matrix(rnorm((p-1)*n),ncol=n)))+ Xbeta
w=matrix(w,ncol=(p-1),byrow=TRUE)
maxw=apply(w,1,max)
y=apply(w,1,indmax)
y=ifelse(maxw < 0,p,y)
return(list(y=y,X=X,beta=beta,sigma=sigma))
}

simout=simmnp(X,p,500,beta,Sigma)

Data1=list(p=p,y=simout$y,X=simout$X)
Mcmc1=list(R=R,keep=1)

out=rmnpGibbs(Data=Data1,Mcmc=Mcmc1)

cat(" Summary of Betadraws ",fill=TRUE)
betatilde=out$betadraw/sqrt(out$sigmadraw[,1])
attributes(betatilde)$class="bayesm.mat"
summary(betatilde,tvalues=beta)

cat(" Summary of Sigmadraws ",fill=TRUE)
sigmadraw=out$sigmadraw/out$sigmadraw[,1]
attributes(sigmadraw)$class="bayesm.var"
summary(sigmadraw,tvalues=as.vector(Sigma[upper.tri(Sigma,diag=TRUE)]))

if(0){
## plotting examples
plot(betatilde,tvalues=beta)
}

```

---

rmultireg

---

*Draw from the Posterior of a Multivariate Regression*


---

## Description

rmultireg draws from the posterior of a Multivariate Regression model with a natural conjugate prior.

## Usage

```
rmultireg(Y, X, Bbar, A, nu, V)
```

## Arguments

Y	n x m matrix of observations on m dep vars
X	n x k matrix of observations on indep vars (supply intercept)
Bbar	k x m matrix of prior mean of regression coefficients
A	k x k Prior precision matrix
nu	d.f. parameter for Sigma
V	m x m pdf location parameter for prior on Sigma

## Details

Model:  $Y = XB + U$ .  $cov(u_i) = Sigma$ .  $B$  is  $k \times m$  matrix of coefficients.  $Sigma$  is  $m \times m$  covariance.

Priors:  $beta$  given  $Sigma \sim N(betabar, Sigma(x)A^{-1})$ .  $betabar = vec(Bbar)$ ;  $beta = vec(B)$   
 $Sigma \sim IW(nu, V)$ .

## Value

A list of the components of a draw from the posterior

B	draw of regression coefficient matrix
Sigma	draw of Sigma

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 2.

<http://www.perossi.org/home/bsm-1>

## Examples

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

set.seed(66)
n=200
m=2
X=cbind(rep(1,n),runif(n))
k=ncol(X)
B=matrix(c(1,2,-1,3),ncol=m)
Sigma=matrix(c(1,.5,.5,1),ncol=m); RSigma=chol(Sigma)
Y=X%B+matrix(rnorm(m*n),ncol=m)%*%RSigma

betabar=rep(0,k*m);Bbar=matrix(betabar,ncol=m)
A=diag(rep(.01,k))
nu=3; V=nu*diag(m)

betadraw=matrix(double(R*k*m),ncol=k*m)
Sigmadraw=matrix(double(R*m*m),ncol=m*m)
for (rep in 1:R)
  {out=rmultireg(Y,X,Bbar,A,nu,V);betadraw[rep,]=out$B
    Sigmadraw[rep,]=out$Sigma}

cat(" Betadraws ",fill=TRUE)
mat=apply(betadraw,2,quantile,probs=c(.01,.05,.5,.95,.99))
mat=rbind(as.vector(B),mat); rownames(mat)[1]="beta"
```

```
print(mat)
cat(" Sigma draws",fill=TRUE)
mat=apply(Sigmadraw,2,quantile,probs=c(.01,.05,.5,.95,.99))
mat=rbind(as.vector(Sigma),mat); rownames(mat)[1]="Sigma"
print(mat)
```

rmvpGibbs

*Gibbs Sampler for Multivariate Probit*

## Description

rmvpGibbs implements the Edwards/Allenby Gibbs Sampler for the multivariate probit model.

## Usage

```
rmvpGibbs(Data, Prior, Mcmc)
```

## Arguments

Data	list(p,y,X)
Prior	list(betabar,A,nu,V) (optional)
Mcmc	list(beta0,sigma0,R,keep) (R required)

## Details

model:

$w_i = X_i \beta + e$ .  $e \sim N(0, \Sigma)$ . note:  $w_i$  is  $p \times 1$ .

$y_{ij} = 1$ , if  $w_{ij} > 0$ , else  $y_{ij} = 0$ .  $j=1, \dots, p$ .

priors:

$\beta \sim N(\text{betabar}, A^{-1})$

$\Sigma \sim \text{IW}(\text{nu}, V)$

to make up X matrix use createX

List arguments contain

- pdimension of multivariate probit
- $X_{n \times p}$  x k Design Matrix
- $y_{n \times p}$  x 1 vector of 0,1 outcomes
- betabar k x 1 prior mean (def: 0)
- A k x k prior precision matrix (def: .01I)
- nu d.f. parm for IWishart prior (def: (p-1) + 3)
- V pds location parm for IWishart prior (def: nu\*I)
- beta0 initial value for beta
- sigma0 initial value for sigma
- R number of MCMC draws
- keep thinning parameter - keep every keepth draw (def: 1)



**Value**

a list containing:

betadraw	R/keep x k array of betadraws
sigmadraw	R/keep x p*p array of sigma draws – each row is in vector form

**Note**

beta and Sigma are not identified. Correlation matrix and the betas divided by the appropriate standard deviation are. See Allenby et al for details or example below.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 4.

<http://www.perossi.org/home/bsm-1>

**See Also**

[rmnpGibbs](#)

**Examples**

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

set.seed(66)
p=3
n=500
beta=c(-2,0,2)
Sigma=matrix(c(1,.5,.5,.5,1,.5,.5,.5,1),ncol=3)
k=length(beta)
I2=diag(rep(1,p)); xadd=rbind(I2)
for(i in 2:n) { xadd=rbind(xadd,I2)}; X=xadd

simmv= function(X,p,n,beta,sigma) {
  w=as.vector(crossprod(chol(sigma),matrix(rnorm(p*n),ncol=n)))+ X%*%beta
  y=ifelse(w<0,0,1)
  return(list(y=y,X=X,beta=beta,sigma=sigma))
}

simout=simmv(X,p,500,beta,Sigma)

Data1=list(p=p,y=simout$y,X=simout$X)
Mcmc1=list(R=R,keep=1)
out=rmvpGibbs(Data=Data1,Mcmc=Mcmc1)

ind=seq(from=0,by=p,length=k)
inda=1:3
ind=ind+inda
cat(" Betadraws ",fill=TRUE)
```

```

betatilde=out$betadraw/sqrt(out$sigmadraw[,ind])
attributes(betatilde)$class="bayesm.mat"
summary(betatilde,tvalues=beta/sqrt(diag(Sigma)))

rdraw=matrix(double((R)*p*p),ncol=p*p)
rdraw=t(apply(out$sigmadraw,1,nmat))
attributes(rdraw)$class="bayesm.var"
tvalue=nmat(as.vector(Sigma))
dim(tvalue)=c(p,p)
tvalue=as.vector(tvalue[upper.tri(tvalue,diag=TRUE)])
cat(" Draws of Correlation Matrix ",fill=TRUE)
summary(rdraw,tvalues=tvalue)

if(0){
plot(betatilde,tvalues=beta/sqrt(diag(Sigma)))
}

```

---

rmvst

---

*Draw from Multivariate Student-t*


---

## Description

rmvst draws from a Multivariate student-t distribution.

## Usage

```
rmvst(nu, mu, root)
```

## Arguments

nu	d.f. parameter
mu	mean vector
root	Upper Tri Cholesky Root of Sigma

## Value

length(mu) draw vector

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch.  
<http://www.perossi.org/home/bsm-1>

**See Also**[lndMvst](#)**Examples**

```
##
set.seed(66)
rmvst(nu=5,mu=c(rep(0,2)),root=chol(matrix(c(2,1,1,2),ncol=2)))
```

rnegbinRw

*MCMC Algorithm for Negative Binomial Regression***Description**

rnegbinRw implements a Random Walk Metropolis Algorithm for the Negative Binomial (NBD) regression model.  $\beta \mid \alpha$  and  $\alpha \mid \beta$  are drawn with two different random walks.

**Usage**

```
rnegbinRw(Data, Prior, Mcmc)
```

**Arguments**

Data	list(y,X)
Prior	list(betabar,A,a,b)
Mcmc	list(R,keep,s\_beta,s\_alpha,beta0)

**Details**

Model:  $y \sim NBD(\text{mean} = \lambda, \text{over} - \text{dispersion} = \alpha)$ .  
 $\lambda = \exp(x'\beta)$

Prior:  $\beta \sim N(\text{betabar}, A^{-1})$

$\alpha \sim \text{Gamma}(a, b)$ .

note: prior mean of  $\alpha = a/b$ , variance =  $a/(b^2)$

list arguments contain:

- y nobs vector of counts (0,1,2,...)
- Xnobs x nvar matrix
- betabar nvar x 1 prior mean (def: 0)
- A nvar x nvar pds prior prec matrix (def: .01I)
- a Gamma prior parm (def: .5)
- b Gamma prior parm (def: .1)
- R number of MCMC draws
- keep MCMC thinning parm: keep every keepth draw (def: 1)
- s\_beta scaling for  $\beta \mid \alpha$  RW inc cov matrix (def:  $2.93/\sqrt{\text{nvar}}$ )
- s\_alpha scaling for  $\alpha \mid \beta$  RW inc cov matrix (def: 2.93)

**Value**

a list containing:

betadraw	R/keep x nvar array of beta draws
alphadraw	R/keep vector of alpha draws
llike	R/keep vector of log-likelihood values evaluated at each draw
acceptrbeta	acceptance rate of the beta draws
acceptralpha	acceptance rate of the alpha draws

**Note**

The NBD regression encompasses Poisson regression in the sense that as alpha goes to infinity the NBD distribution tends toward the Poisson.

For "small" values of alpha, the dependent variable can be extremely variable so that a large number of observations may be required to obtain precise inferences.

**Author(s)**

Sridhar Narayanam & Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby, McCulloch.  
<http://www.perossi.org/home/bsm-1>

**See Also**

[rhierNegbinRw](#)

**Examples**

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=1000} else {R=10}

set.seed(66)
simnegbin =
function(X, beta, alpha) {
# Simulate from the Negative Binomial Regression
lambda = exp(X %*% beta)
y=NULL
for (j in 1:length(lambda))
y = c(y,rnbinom(1,mu = lambda[j],size = alpha))
return(y)
}

nobs = 500
nvar=2          # Number of X variables
alpha = 5
Vbeta = diag(nvar)*0.01

# Construct the regdata (containing X)
simnegbindata = NULL
beta = c(0.6,0.2)
X = cbind(rep(1,nobs),rnorm(nobs,mean=2,sd=0.5))
```

```

simnegbindata = list(y=simnegbin(X,beta,alpha), X=X, beta=beta)

Data1 = simnegbindata
Mcmc1 = list(R=R)

out = rnegbinRw(Data=Data1,Mcmc=Mcmc1)

cat("Summary of alpha/beta draw",fill=TRUE)
summary(out$alphadraw,tvalues=alpha)
summary(out$betadraw,tvalues=beta)

if(0){
## plotting examples
plot(out$betadraw)
}

```

rnmixGibbs

*Gibbs Sampler for Normal Mixtures***Description**

rnmixGibbs implements a Gibbs Sampler for normal mixtures.

**Usage**

```
rnmixGibbs(Data, Prior, Mcmc)
```

**Arguments**

Data	list(y)
Prior	list(Mubar,A,nu,V,a,ncomp) (only ncomp required)
Mcmc	list(R,keep,Loglike) (R required)

**Details**

Model:

$$y_i \sim N(\mu_{ind_i}, \Sigma_{ind_i}).$$

ind  $\sim$  iid multinomial(p). p is a ncomp x 1 vector of probs.

Priors:

$$\mu_j \sim N(\text{mubar}, \Sigma_j(x)A^{-1}). \text{mubar} = \text{vec}(Mubar).$$

$$\Sigma_j \sim \text{IW}(\text{nu}, V).$$

note: this is the natural conjugate prior – a special case of multivariate regression.

$$p \sim \text{Dirchlet}(a).$$

Output of the components is in the form of a list of lists.

compsdraw[[i]] is ith draw – list of ncomp lists.

compsdraw[[i]][[j]] is list of parms for jth normal component.

jcomp=compsdraw[[i]][j]. Then jth comp  $\sim N(jcomp[[1]], \Sigma)$ ,  $\Sigma = t(R)\%*\%R$ ,  $R^{-1} = jcomp[[2]]$ .

List arguments contain:

- y n x k array of data (rows are obs)

- Mubar 1 x k array with prior mean of normal comp means (def: 0)
- A 1 x 1 precision parameter for prior on mean of normal comp (def: .01)
- nu d.f. parameter for prior on Sigma (normal comp cov matrix) (def: k+3)
- V k x k location matrix of IW prior on Sigma (def: nuI)
- a ncomp x 1 vector of Dirichlet prior parms (def: rep(5,ncomp))
- ncomp number of normal components to be included
- R number of MCMC draws
- keep MCMC thinning parm: keep every keepth draw (def: 1)
- LogLike logical flag for compute log-likelihood (def: FALSE)

### Value

nmix	a list containing: probdraw,zdraw,compdraw
ll	vector of log-likelihood values

### Note

more details on contents of nmix:

**probdraw** R/keep x ncomp array of mixture prob draws

**zdraw** R/keep x nobis array of indicators of mixture comp identity for each obs

**compdraw** R/keep lists of lists of comp parm draws

In this model, the component normal parameters are not-identified due to label-switching. However, the fitted mixture of normals density is identified as it is invariant to label-switching. See Allenby et al, chapter 5 for details. Use eMixMargDen or momMix to compute posterior expectation or distribution of various identified parameters.

### Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

### References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 3.

<http://www.perossi.org/home/bsm-1>

### See Also

[rmixture](#), [rmixGibbs](#), [eMixMargDen](#), [momMix](#), [mixDen](#), [mixDenBi](#)

### Examples

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

set.seed(66)
dim=5; k=3 # dimension of simulated data and number of "true" components
sigma = matrix(rep(0.5,dim^2),nrow=dim);diag(sigma)=1
sigfac = c(1,1,1);mufac=c(1,2,3); compsmv=list()
```

```

for(i in 1:k) compsmv[[i]] = list(mu=mufac[i]*1:dim,sigma=sigfac[i]*sigma)
comps = list() # change to "rooti" scale
for(i in 1:k) comps[[i]] = list(mu=compsmv[[i]][[1]],rooti=solve(chol(compsmv[[i]][[2]])))
pvec=(1:k)/sum(1:k)

nobs=500
dm = rmixture(nobs,pvec,comps)

Data1=list(y=dm$x)
ncomp=9
Prior1=list(ncomp=ncomp)
Mcmc1=list(R=R,keep=1)
out=rnmixGibbs(Data=Data1,Prior=Prior1,Mcmc=Mcmc1)

cat("Summary of Normal Mixture Distribution",fill=TRUE)
summary(out)
tmom=momMix(matrix(pvec,nrow=1),list(comps))
mat=rbind(tmom$mu,tmom$sd)
cat(" True Mean/Std Dev",fill=TRUE)
print(mat)

if(0){
##
## plotting examples
##
plot(out$nmix,Data=dm$x)
}

```

rordprobitGibbs

*Gibbs Sampler for Ordered Probit***Description**

rordprobitGibbs implements a Gibbs Sampler for the ordered probit model.

**Usage**

```
rordprobitGibbs(Data, Prior, Mcmc)
```

**Arguments**

Data	list(X, y, k)
Prior	list(betabar, A, dstarbar, Ad)
Mcmc	list(R, keep, s, change, draw)

**Details**

Model:  $z = X\beta + e$ ,  $e \sim N(0, I)$ .  $y=1,...,k$ . cutoff=c( c [1] ,..c [k+1] ).  
 $y=k$ , if  $c [k] \leq z < c [k+1]$  .

Prior:  $\beta \sim N(\text{betabar}, A^{-1})$ .  $dstar \sim N(\text{dstarbar}, Ad^{-1})$ .

List arguments contain

$X$   $n \times nvar$  Design Matrix  
 $y$   $n \times 1$  vector of observations, (1,...,k)  
 $k$  the largest possible value of  $y$   
 $betabar$   $nvar \times 1$  prior mean (def: 0)  
 $A$   $nvar \times nvar$  prior precision matrix (def: .01I)  
 $dstarbar$   $ndstar \times 1$  prior mean,  $ndstar=k-2$  (def: 0)  
 $Ad$   $ndstar \times ndstar$  prior precision matrix (def:I)  
 $s$  scaling parm for RW Metropolis (def:  $2.93/\sqrt{nvar}$ )  
 $R$  number of MCMC draws  
 $keep$  thinning parameter - keep every  $keepth$  draw (def: 1)

### Value

$betadraw$	$R/keep \times k$ matrix of betadraws
$cutdraw$	$R/keep \times (k-1)$ matrix of cutdraws
$dstardraw$	$R/keep \times (k-2)$ matrix of dstardraws
$accept$	a value of acceptance rate in RW Metropolis

### Note

set  $c[1]=-100$ .  $c[k+1]=100$ .  $c[2]$  is set to 0 for identification.

The relationship between cut-offs and  $dstar$  is

$c[3] = \exp(dstar[1])$ ,  $c[4]=c[3]+\exp(dstar[2])$ ,...,  $c[k] = c[k-1] + \exp(dstar[k-2])$

Be careful in assessing prior parameter,  $Ad$ . .1 is too small for many applications.

### Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

### References

*Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch  
<http://www.perossi.org/home/bsm-1>

### See Also

[rbprobitGibbs](#)

### Examples

```
##
## rordprobitGibbs example
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}

## simulate data for ordered probit model

simordprobit=function(X, betas, cutoff){
  z = X%*%betas + rnorm(nobs)
```



```

    y = cut(z, br = cutoff, right=TRUE, include.lowest = TRUE, labels = FALSE)
    return(list(y = y, X = X, k=(length(cutoff)-1), betas= betas, cutoff=cutoff ))
  }

  set.seed(66)
  nobs=300
  X=cbind(rep(1,nobs),runif(nobs, min=0, max=5),runif(nobs,min=0, max=5))
  k=5
  betas=c(0.5, 1, -0.5)
  cutoff=c(-100, 0, 1.0, 1.8, 3.2, 100)
  simout=simordprobit(X, betas, cutoff)
  Data=list(X=simout$X,y=simout$y, k=k)

## set Mcmc for ordered probit model

Mcmc=list(R=R)
out=rordprobitGibbs(Data=Data,Mcmc=Mcmc)

cat(" ", fill=TRUE)
cat("acceptance rate= ",accept=out$accept,fill=TRUE)

## outputs of betadraw and cut-off draws

cat(" Summary of betadraws",fill=TRUE)
summary(out$betadraw,tvalues=betas)
cat(" Summary of cut-off draws",fill=TRUE)
summary(out$cutdraw,tvalues=cutoff[2:k])

if(0){
## plotting examples
plot(out$cutdraw)
}

```

rscaleUsage

*MCMC Algorithm for Multivariate Ordinal Data with Scale Usage Heterogeneity.*

## Description

rscaleUsage implements an MCMC algorithm for multivariate ordinal data with scale usage heterogeneity.

## Usage

```
rscaleUsage(Data,Prior, Mcmc)
```

## Arguments

Data	list(k,x)
Prior	list(nu,V,mubar,Am,gsigma,gl11,gl22,gl12,Lambdanu,LambdaV,ge) (optional)
Mcmc	list(R,keep,ndghk,printevery,e,y,mu,Sigma,sigma,tau,Lambda) (optional)

## Details

Model:  $n = \text{nrow}(x)$  individuals respond to  $m = \text{ncol}(x)$  questions. all questions are on a scale  $1, \dots, k$ . for respondent  $i$  and question  $j$ ,

$$x_{ij} = d, \text{ if } c_{d-1} \leq y_{ij} \leq c_d.$$

$$d = 1, \dots, k. \quad c_d = a + bd + ed^2.$$

$$y_i = \mu + \tau_i * \text{iota} + \sigma_i * z_i. \quad z_i \sim N(0, \text{Sigma}).$$

Priors:

$$(\tau_i, \ln(\sigma_i)) \sim N(\phi, \text{Lamda}). \quad \phi = (0, \text{lambda}_{22}).$$

$$\mu \sim N(\text{mubar}, A m^{-1}).$$

$$\text{Sigma} \sim \text{IW}(\text{nu}, V).$$

$$\text{Lambda} \sim \text{IW}(\text{Lambdanu}, \text{LambdaV}).$$

$$e \sim \text{unif on a grid}.$$

## Value

a list containing:

Sigmadraw	R/keep x m*m array of Sigma draws
mudraw	R/keep x m array of mu draws
taudraw	R/keep x n array of tau draws
sigmadraw	R/keep x n array of sigma draws
Lambdadraw	R/keep x 4 array of Lamda draws
edraw	R/keep x 1 array of e draws

## Warning

$\tau_i, \sigma_i$  are identified from the scale usage patterns in the  $m$  questions asked per respondent ( $\# \text{ cols of } x$ ). Do not attempt to use this on data sets with only a small number of total questions!

## Note

It is **highly** recommended that the user choose the default settings. This means not specifying the argument `Prior` and setting `R` in `Mcmc` and `Data` only. If you wish to change prior settings and/or the grids used, please read the case study in Allenby et al carefully.

## Author(s)

Rob McCulloch and Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby, and McCulloch, Case Study on Scale Usage Heterogeneity.

<http://www.perossi.org/home/bsm-1>

## Examples

```
##
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=1000} else {R=1}
{
  data(customerSat)
  surveydat = list(k=10,x=as.matrix(customerSat))

  Mcmc1 = list(R=R)
  set.seed(66)
  out=rscaleUsage(Data=surveydat,Mcmc=Mcmc1)

  summary(out$mudraw)
}
```

rsurGibbs

*Gibbs Sampler for Seemingly Unrelated Regressions (SUR)*

## Description

rsurGibbs implements a Gibbs Sampler to draw from the posterior of the Seemingly Unrelated Regression (SUR) Model of Zellner

## Usage

```
rsurGibbs(Data, Prior, Mcmc)
```

## Arguments

Data	list(regdata)
Prior	list(betabar,A, nu, V)
Mcmc	list(R,keep)

## Details

Model:  $y_i = X_i \beta_i + e_i$ ,  $i=1, \dots, m$ .  $m$  regressions.  
 $(e(1,k), \dots, e(m,k)) \sim N(0, \text{Sigma})$ .  $k=1, \dots, \text{nobs}$ .

We can also write as the stacked model:

$y = X\beta + e$  where  $y$  is a  $\text{nobs} \times m$  long vector and  $k=\text{length}(\beta)=\text{sum}(\text{length}(\beta_i))$ .

Note: we must have the same number of observations in each equation but we can have different numbers of  $X$  variables

Priors:  $\beta \sim N(\text{betabar}, A^{-1})$ .  $\text{Sigma} \sim IW(\text{nu}, V)$ .

List arguments contain

- regdatalist of lists, regdata[[i]]=list(y=yi,X=Xi)
- betabark x 1 prior mean (def: 0)
- Ak x k prior precision matrix (def: .01I)
- nu d.f. parm for Inverted Wishart prior (def: m+3)
- V scale parm for Inverted Wishart prior (def: nu\*I)
- R number of MCMC draws
- keep thinning parameter - keep every keepth draw

**Value**

list of MCMC draws

betadraw            R x k array of betadraws

Sigmadraw          R x (m\*m) array of Sigma draws

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 3.

<http://www.perossi.org/home/bsm-1>

**See Also**

[rmultireg](#)

**Examples**

```
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=1000} else {R=10}
##
## simulate data from SUR
set.seed(66)
beta1=c(1,2)
beta2=c(1,-1,-2)
nobs=100
nreg=2
iota=c(rep(1,nobs))
X1=cbind(iota,runif(nobs))
X2=cbind(iota,runif(nobs),runif(nobs))
Sigma=matrix(c(.5,.2,.2,.5),ncol=2)
U=chol(Sigma)
E=matrix(rnorm(2*nobs),ncol=2)%*%U
y1=X1%*%beta1+E[,1]
y2=X2%*%beta2+E[,2]
##
## run Gibbs Sampler
regdata=NULL
regdata[[1]]=list(y=y1,X=X1)
regdata[[2]]=list(y=y2,X=X2)

Mcmc1=list(R=R)

out=rsurGibbs(Data=list(regdata=regdata),Mcmc=Mcmc1)

cat("Summary of beta draws",fill=TRUE)
summary(out$betadraw,tvalues=c(beta1,beta2))
cat("Summary of Sigmadrws",fill=TRUE)
summary(out$Sigmadraw,tvalues=as.vector(Sigma[upper.tri(Sigma,diag=TRUE)]))

if(0){
plot(out$betadraw,tvalues=c(beta1,beta2))
}
```

---

rtrun*Draw from Truncated Univariate Normal*

---

## Description

rtrun draws from a truncated univariate normal distribution

## Usage

```
rtrun(mu, sigma, a, b)
```

## Arguments

mu	mean
sigma	sd
a	lower bound
b	upper bound

## Details

Note that due to the vectorization of the `rnorm`, `qnorm` commands in R, all arguments can be vectors of equal length. This makes the inverse CDF method the most efficient to use in R.

## Value

draw (possibly a vector)

## Warning

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 2.

<http://www.perossi.org/home/bsm-1>

## Examples

```
##  
set.seed(66)  
rtrun(mu=c(rep(0,10)),sigma=c(rep(1,10)),a=c(rep(0,10)),b=c(rep(2,10)))
```

runireg

*IID Sampler for Univariate Regression***Description**

runireg implements an iid sampler to draw from the posterior of a univariate regression with a conjugate prior.

**Usage**

```
runireg(Data, Prior, Mcmc)
```

**Arguments**

Data	list(y,X)
Prior	list(betabar,A, nu, ssq)
Mcmc	list(R,keep)

**Details**

Model:  $y = X\beta + e$ .  $e \sim N(0, \sigma^2)$ .

Priors:  $\beta \sim N(\text{betabar}, \sigma^2 A^{-1})$ .  $\sigma^2 \sim (nu * ssq) / \text{chisq}_{nu}$ . List arguments contain

- $X$  n x k Design Matrix
- $y$  n x 1 vector of observations
- betabar k x 1 prior mean (def: 0)
- $A$  k x k prior precision matrix (def: .01I)
- nu d.f. parm for Inverted Chi-square prior (def: 3)
- ssq scale parm for Inverted Chi-square prior (def: var(y))
- R number of draws
- keep thinning parameter - keep every keepth draw

**Value**

list of iid draws

betadraw	R x k array of betadraws
sigmasqdraw	R vector of sigma-sq draws

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 2.

<http://www.perossi.org/home/bsm-1>

**See Also**[runiregGibbs](#)**Examples**

```

if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=2000} else {R=10}
set.seed(66)
n=200
X=cbind(rep(1,n),runif(n)); beta=c(1,2); sigsq=.25
y=X%%beta+rnorm(n,sd=sqrt(sigsq))

out=runireg(Data=list(y=y,X=X),Mcmc=list(R=R))

cat("Summary of beta/sigma-sq draws",fill=TRUE)
summary(out$betadraw,tvalues=beta)
summary(out$sigmasqdraw,tvalues=sigsq)

if(0){
## plotting examples
plot(out$betadraw)
}

```

runiregGibbs

*Gibbs Sampler for Univariate Regression***Description**

runiregGibbs implements a Gibbs Sampler to draw from posterior of a univariate regression with a conditionally conjugate prior.

**Usage**

```
runiregGibbs(Data, Prior, Mcmc)
```

**Arguments**

Data	list(y,X)
Prior	list(betabar,A, nu, ssq)
Mcmc	list(sigmasq,R,keep)

**Details**

Model:  $y = X\beta + e$ .  $e \sim N(0, \text{sigmasq})$ .

Priors:  $\beta \sim N(\text{betabar}, A^{-1})$ .  $\text{sigmasq} \sim (nu * ssq) / \text{chisq}_{nu}$ . List arguments contain

- $X$  n x k Design Matrix
- $y$  n x 1 vector of observations
- $\text{betabar}$  k x 1 prior mean (def: 0)
- $A$  k x k prior precision matrix (def: .01I)

- nu d.f. parm for Inverted Chi-square prior (def: 3)
- ssq scale parm for Inverted Chi-square prior (def:var(y))
- R number of MCMC draws
- keep thinning parameter - keep every keepth draw

### Value

list of MCMC draws

betadraw            R x k array of betadraws

sigmasqdraw       R vector of sigma-sq draws

### Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

### References

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 3.

<http://www.perossi.org/home/bsm-1>

### See Also

[runireg](#)

### Examples

```
if(nchar(Sys.getenv("LONG_TEST")) != 0) {R=1000} else {R=10}
set.seed(66)
n=100
X=cbind(rep(1,n),runif(n)); beta=c(1,2); sigsq=.25
y=X*%beta+rnorm(n,sd=sqrt(sigsq))

Data1=list(y=y,X=X); Mcmc1=list(R=R)

out=runiregGibbs(Data=Data1,Mcmc=Mcmc1)

cat("Summary of beta and Sigma draws",fill=TRUE)
summary(out$betadraw,tvalues=beta)
summary(out$sigmasqdraw,tvalues=sigsq)

if(0){
## plotting examples
plot(out$betadraw)
}
```



rwishart

*Draw from Wishart and Inverted Wishart Distribution***Description**

rwishart draws from the Wishart and Inverted Wishart distributions.

**Usage**

```
rwishart(nu, V)
```

**Arguments**

nu	d.f. parameter
V	pds location matrix

**Details**

In the parameterization used here,  $W \sim W(nu, V)$ ,  $E[W] = nuV$ .

If you want to use an Inverted Wishart prior, you *must invert the location matrix* before calling rwishart, e.g.

$Sigma \sim IW(nu, V)$ ;  $Sigma^{-1} \sim W(nu, V^{-1})$ .

**Value**

W	Wishart draw
IW	Inverted Wishart draw
C	Upper tri root of W
CI	$inv(C)$ , $W^{-1} = CICI'$

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 2.

<http://www.perossi.org/home/bsm-1>

**Examples**

```
##
set.seed(66)
rwishart(5, diag(3))$IW
```

---

Scotch

---

*Survey Data on Brands of Scotch Consumed*


---

**Description**

from Simmons Survey. Brands used in last year for those respondents who report consuming scotch.

**Usage**

```
data(Scotch)
```

**Format**

A data frame with 2218 observations on the following 21 variables. All variables are coded 1 if consumed in last year, 0 if not.

Chivas.Regal a numeric vector  
 Dewar.s.White.Label a numeric vector  
 Johnnie.Walker.Black.Label a numeric vector  
 J...B a numeric vector  
 Johnnie.Walker.Red.Label a numeric vector  
 Other.Brands a numeric vector  
 Glenlivet a numeric vector  
 Cutty.Sark a numeric vector  
 Glenfiddich a numeric vector  
 Pinch..Haig. a numeric vector  
 Clan.MacGregor a numeric vector  
 Ballantine a numeric vector  
 Macallan a numeric vector  
 Passport a numeric vector  
 Black...White a numeric vector  
 Scoresby.Rare a numeric vector  
 Grants a numeric vector  
 Ushers a numeric vector  
 White.Horse a numeric vector  
 Knockando a numeric vector  
 the.Singleton a numeric vector

**Source**

Edwards, Y. and G. Allenby (2003), "Multivariate Analysis of Multiple Response Data," *JMR* 40, 321-334.

**References**

Chapter 4, *Bayesian Statistics and Marketing* by Rossi et al.  
<http://www.perossi.org/home/bsm-1>

**Examples**

```

data(Scotch)
cat(" Frequencies of Brands", fill=TRUE)
mat=apply(as.matrix(Scotch),2,mean)
print(mat)
##
## use Scotch data to run Multivariate Probit Model
##
if(0){
##

y=as.matrix(Scotch)
p=ncol(y); n=nrow(y)
dimnames(y)=NULL
y=as.vector(t(y))
y=as.integer(y)
I_p=diag(p)
X=rep(I_p,n)
X=matrix(X,nrow=p)
X=t(X)

R=2000
Data=list(p=p,X=X,y=y)
Mcmc=list(R=R)
set.seed(66)
out=rmvpGibbs(Data=Data,Mcmc=Mcmc)

ind=(0:(p-1))*p + (1:p)
cat(" Betadraws ",fill=TRUE)
mat=apply(out$betadraw/sqrt(out$sigmadraw[,ind]),2,quantile,probs=c(.01,.05,.5,.95,.99))
attributes(mat)$class="bayesm.mat"
summary(mat)
rdraw=matrix(double((R)*p*p),ncol=p*p)
rdraw=t(apply(out$sigmadraw,1,nmat))
attributes(rdraw)$class="bayesm.var"
cat(" Draws of Correlation Matrix ",fill=TRUE)
summary(rdraw)

}

```

---

simnhlogit

---

*Simulate from Non-homothetic Logit Model*


---

**Description**

simnhlogit simulates from the non-homothetic logit model

**Usage**

```
simnhlogit(theta, lnprices, Xexpend)
```

**Arguments**

theta	coefficient vector
lnprices	n x p array of prices
Xexpend	n x k array of values of expenditure variables

**Details**

For detail on parameterization, see `llnhlogit`.

**Value**

a list containing:

y	n x 1 vector of multinomial outcomes (1, ..., p)
Xexpend	expenditure variables
lnprices	price array
theta	coefficients
prob	n x p array of choice probabilities

**Warning**

This routine is a utility routine that does **not** check the input arguments for proper dimensions and type.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**References**

For further discussion, see *Bayesian Statistics and Marketing* by Rossi, Allenby and McCulloch, Chapter 4.

<http://www.perossi.org/home/bsm-1>

**See Also**

[llnhlogit](#)

---

summary.bayesm.mat	<i>Summarize Mcmc Parameter Draws</i>
--------------------	---------------------------------------

---

**Description**

summary.bayesm.mat is an S3 method to summarize marginal distributions given an array of draws

**Usage**

```
## S3 method for class 'bayesm.mat'
summary(object, names, burnin = trunc(0.1 * nrow(X)), tvalues, QUANTILES = TRUE, TRAILER = TRUE, ...)
```

**Arguments**

object	object (hereafter X) is an array of draws, usually an object of class "bayesm.mat"
names	optional character vector of names for the columns of X
burnin	number of draws to burn-in, def: .1*nrow(X)
tvalues	optional vector of "true" values for use in simulation examples
QUANTILES	logical for should quantiles be displayed, def: TRUE
TRAILER	logical for should a trailer be displayed, def: TRUE
...	optional arguments for generic function

**Details**

Typically, `summary.bayesm.nmix` will be invoked by a call to the generic summary function as in `summary(object)` where `object` is of class `bayesm.mat`. Mean, Std Dev, Numerical Standard error (of estimate of posterior mean), relative numerical efficiency (see `numEff`) and effective sample size are displayed. If `QUANTILES=TRUE`, quantiles of marginal distributions in the columns of X are displayed.

`summary.bayesm.mat` is also exported for direct use as a standard function, as in `summary.bayesm.mat(matrix)`. `summary.bayesm.mat(matrix)` returns (invisibly) the array of the various summary statistics for further use. To assess this array use `stats=summary(Drawmat)`.

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**See Also**

[summary.bayesm.var](#), [summary.bayesm.nmix](#)

**Examples**

```
##
## not run
# out=rmpGibbs(Data,Prior,Mcmc)
# summary(out$betadraw)
#
```

---

summary.bayesm.nmix	<i>Summarize Draws of Normal Mixture Components</i>
---------------------	---

---

**Description**

`summary.bayesm.nmix` is an S3 method to display summaries of the distribution implied by draws of Normal Mixture Components. Posterior means and Variance-Covariance matrices are displayed.

Note: 1st and 2nd moments may not be very interpretable for mixtures of normals. This summary function can take a minute or so. The current implementation is not efficient.

**Usage**

```
## S3 method for class 'bayesm.nmix'
summary(object, names, burnin = trunc(0.1 * nrow(probdraw)), ...)
```

**Arguments**

object	an object of class "bayesm.nmix" – a list of lists of draws
names	optional character vector of names fo reach dimension of the density
burnin	number of draws to burn-in, def: .1*nrow(probdraw)
...	parms to send to summary

**Details**

an object of class "bayesm.nmix" is a list of three components:

**probdraw** a matrix of R/keep rows by dim of normal mix of mixture prob draws

**second comp** not used

**compdraw** list of list of lists with draws of mixture comp parms

**Author(s)**

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

**See Also**

[summary.bayesm.mat](#), [summary.bayesm.var](#)

**Examples**

```
##
## not run
# out=rnmix(Data,Prior,Mcmc)
# summary(out)
#
```

---

summary.bayesm.var	<i>Summarize Draws of Var-Cov Matrices</i>
--------------------	--

---

**Description**

summary.bayesm.var is an S3 method to summarize marginal distributions given an array of draws

**Usage**

```
## S3 method for class 'bayesm.var'
summary(object, names, burnin = trunc(0.1 * nrow(Vard)), tvalues, QUANTILES = FALSE , ...)
```

## Arguments

object	object (hereafter, Vard) is an array of draws of a covariance matrix
names	optional character vector of names for the columns of Vard
burnin	number of draws to burn-in, def: .1*nrow(Vard)
tvalues	optional vector of "true" values for use in simulation examples
QUANTILES	logical for should quantiles be displayed, def: TRUE
...	optional arguments for generic function

## Details

Typically, `summary.bayesm.var` will be invoked by a call to the generic summary function as in `summary(object)` where object is of class `bayesm.var`. Mean, Std Dev, Numerical Standard error (of estimate of posterior mean), relative numerical efficiency (see `numEff`) and effective sample size are displayed. If `QUANTILES=TRUE`, quantiles of marginal distributions in the columns of Vard are displayed.

Vard is an array of draws of a covariance matrix stored as vectors. Each row is a different draw. The posterior mean of the vector of standard deviations and the correlation matrix are also displayed

## Author(s)

Peter Rossi, Anderson School, UCLA, <perossichi@gmail.com>.

## See Also

[summary.bayesm.mat](#), [summary.bayesm.nmix](#)

## Examples

```
##
## not run
# out=rmpGibbs(Data,Prior,Mcmc)
# summary(out$sigmadraw)
#
```

---

tuna

*Data on Canned Tuna Sales*

---

## Description

Volume of canned tuna sales as well as a measure of display activity, log price and log wholesale price. Weekly data aggregated to the chain level. This data is extracted from the Dominick's Finer Foods database maintained by the University of Chicago <http://research.chicagogsb.edu/marketing/databases/dominicks/dataset.aspx>. Brands are seven of the top 10 UPCs in the canned tuna product category.

## Usage

```
data(tuna)
```

### Format

A data frame with 338 observations on the following 30 variables.

WEEK a numeric vector

MOVE1 unit sales of Star Kist 6 oz.

MOVE2 unit sales of Chicken of the Sea 6 oz.

MOVE3 unit sales of Bumble Bee Solid 6.12 oz.

MOVE4 unit sales of Bumble Bee Chunk 6.12 oz.

MOVE5 unit sales of Geisha 6 oz.

MOVE6 unit sales of Bumble Bee Large Cans.

MOVE7 unit sales of HH Chunk Lite 6.5 oz.

NSALE1 a measure of display activity of Star Kist 6 oz.

NSALE2 a measure of display activity of Chicken of the Sea 6 oz.

NSALE3 a measure of display activity of Bumble Bee Solid 6.12 oz.

NSALE4 a measure of display activity of Bumble Bee Chunk 6.12 oz.

NSALE5 a measure of display activity of Geisha 6 oz.

NSALE6 a measure of display activity of Bumble Bee Large Cans.

NSALE7 a measure of display activity of HH Chunk Lite 6.5 oz.

LPRICE1 log of price of Star Kist 6 oz.

LPRICE2 log of price of Chicken of the Sea 6 oz.

LPRICE3 log of price of Bumble Bee Solid 6.12 oz.

LPRICE4 log of price of Bumble Bee Chunk 6.12 oz.

LPRICE5 log of price of Geisha 6 oz.

LPRICE6 log of price of Bumble Bee Large Cans.

LPRICE7 log of price of HH Chunk Lite 6.5 oz.

LWHPRIC1 log of wholesale price of Star Kist 6 oz.

LWHPRIC2 log of wholesale price of Chicken of the Sea 6 oz.

LWHPRIC3 log of wholesale price of Bumble Bee Solid 6.12 oz.

LWHPRIC4 log of wholesale price of Bumble Bee Chunk 6.12 oz.

LWHPRIC5 log of wholesale price of Geisha 6 oz.

LWHPRIC6 log of wholesale price of Bumble Bee Large Cans.

LWHPRIC7 log of wholesale price of HH Chunk Lite 6.5 oz.

FULLCUST total customers visits

### Source

Chevalier, A. Judith, Anil K. Kashyap and Peter E. Rossi (2003), "Why Don't Prices Rise During Periods of Peak Demand? Evidence from Scanner Data," *The American Economic Review*, 93(1), 15-37.

### References

Chapter 7, *Bayesian Statistics and Marketing* by Rossi et al.

<http://www.perossi.org/home/bsm-1>



**Examples**

```

data(tuna)
cat(" Quantiles of sales",fill=TRUE)
mat=apply(as.matrix(tuna[,2:5]),2,quantile)
print(mat)

##
## example of processing for use with rivGibbs
##
if(0)
{
  data(tuna)
  t = dim(tuna)[1]
  customers = tuna[,30]
  sales = tuna[,2:8]
  lnprice = tuna[,16:22]
  lnwhPrice= tuna[,23:29]
  share=sales/mean(customers)
  shareout=as.vector(1-rowSums(share))
  lnprob=log(share/shareout)

  # create w matrix

  I1=as.matrix(rep(1, t))
  I0=as.matrix(rep(0, t))
  intercept=rep(I1, 4)
  brand1=rbind(I1, I0, I0, I0)
  brand2=rbind(I0, I1, I0, I0)
  brand3=rbind(I0, I0, I1, I0)
  w=cbind(intercept, brand1, brand2, brand3)

  ## choose brand 1 to 4

  y=as.vector(as.matrix(lnprob[,1:4]))
  X=as.vector(as.matrix(lnprice[,1:4]))
  lnwhPrice=as.vector(as.matrix (lnwhPrice[1:4]))
  z=cbind(w, lnwhPrice)

  Data=list(z=z, w=w, x=X, y=y)
  Mcmc=list(R=R, keep=1)
  set.seed(66)
  out=rivGibbs(Data=Data,Mcmc=Mcmc)

  cat(" betadraws ",fill=TRUE)
  summary(out$betadraw)

  if(0){
  ## plotting examples
  plot(out$betadraw)
  }
  }

```

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